



Answer the following questions.

Q1: [10]

Determine the survival, density, hazard rate, mean excess loss and limited expected value functions for the age at death model that defined as:

$$F(x) = \begin{cases} 0, & x < 0 \\ 0.01x, & 0 \leq x < 100 \\ 1, & x \geq 100 \end{cases}$$

Q2: [5+5]

(a) Show that the gamma distribution is a member of the linear exponential family, then derive the mean and variance of the gamma distribution.

(b) Let  $\Lambda$  have an exponential distribution and let  $X|\Lambda$  have a Weibull distribution with conditional survival function  $S_{X|\Lambda}(x|\lambda) = e^{-\lambda x^\gamma}$ ,  $x \geq 0$ . Determine the unconditional or marginal distribution of  $X$ .

Q3: [5+5]

(a) Seven losses are observed as 27, 82, 115, 126, 155, 161 and 243. Determine the maximum likelihood estimate of the parameter  $\theta$  for an exponential distribution, and for a gamma distribution with  $\alpha=2$ . Also, find the value of the log-likelihood function in each case.

(b) Let  $X$  have a Pareto distribution with parameters  $\alpha$  and  $\theta$ . Let  $Y = \ln(1 + X/\theta)$ . Determine the distribution of  $Y$  and its parameters (give its name).

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## The Model Answer

Q1: [10]

(i)  $S(x) = 1 - 0.01x$ ,  $0 \leq x < 100$

(ii)  $f(x) = F'(x) = -S'(x) = 0.01$

(iii)

$$\begin{aligned}\therefore h(x) &= \frac{f(x)}{S(x)} \\ \therefore h(x) &= \frac{0.01}{1 - 0.01x}\end{aligned}$$

(v)

$$\begin{aligned}\therefore e_X(d) &= \frac{\int_d^\infty S(x)dx}{S(d)} \\ &= \frac{\int_d^{100} (1 - 0.01x)dx}{1 - 0.01d} \\ &= \frac{-1}{0.01(1 - 0.01d)} \left[ \frac{(1 - 0.01x)^2}{2} \right]_d^{100} \\ \therefore e_X(d) &= \frac{1 - 0.01d}{0.02} = \frac{100 - d}{2}\end{aligned}$$

(vi)

$$\begin{aligned}\therefore E(X \wedge u) &= \int_{-\infty}^u xf(x)dx + u[1 - F(u)] \\ &= \int_0^u x(0.01)dx + u(1 - 0.01u) \\ &= 0.005u^2 + u - 0.01u^2\end{aligned}$$

$$\therefore E(X \wedge u) = u(1 - 0.005u)$$

or

$$\begin{aligned}
E(X \wedge u) &= \int_0^u S(x) dx \\
&= \int_0^u (1 - 0.01x) dx \\
&= \left[ x - \frac{0.01x^2}{2} \right]_0^u \\
\therefore E(X \wedge u) &= u - 0.005u^2
\end{aligned}$$

or

By using the following formula

$$E(X \wedge u) = E(X) - e(u)S(u)$$

$$\begin{aligned}
\therefore E(X) &= \int_0^u x(0.01) dx = \int_0^{100} x(0.01) dx \\
&= 0.01 \left[ \frac{x^2}{2} \right]_0^{100}
\end{aligned}$$

$$\therefore E(x) = 50$$

$$\text{Also, } e(u) = \frac{100-u}{2} \text{ and } S(u) = 1 - 0.01u$$

$$\begin{aligned}
\therefore E(X \wedge u) &= E(X) - e(u)S(u) \\
&= 50 - \left(\frac{100-u}{2}\right)(1 - 0.01u) \\
&= 50 - \frac{0.01}{2}(100-u)^2
\end{aligned}$$

$$\therefore E(X \wedge u) = u - 0.005u^2$$

**Q2: [5+5]**

(a)

For  $X \sim \text{gamma}(\alpha, \theta)$

$$\Rightarrow f(x; \theta) = \frac{\theta^{-\alpha} x^{\alpha-1} e^{-x/\theta}}{\Gamma(\alpha)}$$

$$\text{Clearly, } f(x; \theta) = \frac{p(x)e^{r(\theta)x}}{q(\theta)}$$

$$= \frac{[x^{\alpha-1} / \Gamma(\alpha)] \cdot e^{-\frac{1}{\theta}x}}{\theta^\alpha}$$

Where,  $r(\theta) = -1/\theta$ ,  $q(\theta) = \theta^\alpha$  and  $p(x) = x^{\alpha-1} / \Gamma(\alpha)$

∴ The gamma distribution is a member of the linear exponential family.

$$\begin{aligned} \therefore \text{The mean, } E(X) = \mu(\theta) &= \frac{q'(\theta)}{r'(\theta)q(\theta)} \\ &= \frac{\alpha\theta^{\alpha-1}}{1/\theta^2 \cdot \theta^\alpha} = \alpha\theta \end{aligned}$$

and the variance,  $Var(X) = v(\theta)$

$$\begin{aligned} &= \frac{\mu'(\theta)}{r'(\theta)} \\ &= \frac{\alpha}{1/\theta^2} = \alpha\theta^2 \end{aligned}$$

(b)

$$\therefore X|\Lambda \sim \text{weibull}(\lambda, \gamma), S_{X|\Lambda}(x|\lambda) = e^{-\lambda x^\gamma}$$

$$\therefore A(x) = x^\gamma$$

$$\therefore \Lambda \sim \exp(\theta)$$

$$\therefore M_\Lambda(z) = (1 - \theta z)^{-1}$$

$$\begin{aligned} \therefore S_X(x) &= E[e^{-\Lambda A(x)}] \\ &= M_\Lambda[-A(x)] \end{aligned}$$

$$\therefore S_X(x) = M_\Lambda[-x^\gamma]$$

$$\therefore S_X(x) = (1 + \theta x^\gamma)^{-1}$$

Which is a loglogistic distribution with the usual parameter  $\theta$  replaced by  $\theta^{\frac{1}{\gamma}}$ .

**Q3: [5+5]**

For exponential distribution, the likelihood function is

$$\begin{aligned}
L(\theta) &= f(27)f(82)f(115)f(126)f(155)f(161)f(243) \\
&= \theta^{-1}e^{-27/\theta}\theta^{-1}e^{-82/\theta}\theta^{-1}e^{-115/\theta}\theta^{-1}e^{-126/\theta}\theta^{-1}e^{-155/\theta}\theta^{-1}e^{-161/\theta}\theta^{-1}e^{-243/\theta} \\
&= \theta^{-7}e^{-909/\theta}
\end{aligned}$$

$$\therefore l(\theta) = -7\ln\theta - 909\theta^{-1}$$

which is known as log-likelihood function, to get the likelihood estimate of the parameter  $\theta$

$$\text{Set } l'(\theta) = 0$$

$$\Rightarrow -7\theta^{-1} + 909\theta^{-2} = 0$$

$\therefore \hat{\theta} = 129.85714$  which is the MLE of the mean of an exponential model.

$$\therefore l(\hat{\theta}) = -41.065$$

For a gamma distribution with  $\alpha=2$ ,  $f(x_j|\theta) = x_j\theta^{-2}e^{-x_j/\theta}$ ,  $j=1,2,\dots,7$

$$\begin{aligned}
\therefore l(\theta) &= \sum_{j=1}^n \ln(f(x_j|\theta)) = \sum_{j=1}^7 \ln(x_j\theta^{-2}e^{-x_j/\theta}) \\
&= \sum_{j=1}^7 \ln x_j - 14\ln\theta - \theta^{-1} \sum_{j=1}^7 x_j
\end{aligned}$$

$$\text{Set } l'(\theta) = 0$$

$$\Rightarrow -14\theta^{-1} + \theta^{-2} \sum_{j=1}^7 x_j = 0$$

$$-14\theta^{-1} + 7\theta^{-2} \bar{x} = 0 \quad (\times \frac{\theta^2}{14})$$

$$\therefore \hat{\theta} = \frac{\bar{x}}{2}$$

$$\therefore \hat{\theta} = 64.9286$$

$$\Rightarrow l(\hat{\theta}) = \sum_{j=1}^7 \ln x_j - 14\ln\hat{\theta} - 14$$

$$\therefore l(\hat{\theta}) = -39.5244$$

(b)

$\therefore X \sim \text{Pareto}(\alpha, \theta)$

$$\therefore F_X(x) = 1 - \left( \frac{\theta}{x + \theta} \right)^\alpha$$

For  $Y = \ln(1 + X / \theta)$ ,  $F_Y(y) = \Pr(Y \leq y)$

$$F_Y(y) = \Pr[\ln(1 + X / \theta) \leq y]$$

$$\therefore F_Y(y) = \Pr[1 + X / \theta \leq e^y]$$

$$= \Pr[X \leq \theta(e^y - 1)]$$

$$= 1 - \left[ \frac{\theta}{\theta(e^y - 1) + \theta} \right]^\alpha$$

$$\therefore F_Y(y) = 1 - \left( \frac{1}{e^y} \right)^\alpha$$

$$= 1 - e^{-\alpha y},$$

which is the distribution function of the exponential distribution with parameter  $1 / \alpha$ .

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