

Lecture (15)

• Sec 5.2.5 Frailty Models p. (65)

let $\Lambda > 0$ be frailty random variable and define the conditional hazard rate

$$h_{X|\Lambda}(x|\lambda) = \lambda a(x),$$

$a(x)$ is known function for a certain application.

\Rightarrow the conditional survival fn of $X|\Lambda$ is therefore

$$\begin{aligned} S_{X|\Lambda}(x|\lambda) &= e^{-\int_0^x h_{X|\Lambda}(t|\lambda) dt} \\ &= e^{-\lambda \int_0^x a(t) dt} \end{aligned}$$

$$\therefore S_{X|\Lambda}(x|\lambda) = e^{-\lambda A(x)},$$

where $A(x) = \int_0^x a(t) dt$

* In order to specify the mixture distribution (i.e. the marginal distn of X), we define the moment generating fn of the frailty random variable Λ to be

$$M_{\Lambda}(z) = E(e^{z\Lambda}).$$

• Then, the marginal survival function is

$$S_X(x) = E[e^{-\Lambda A(x)}] = M_{\Lambda}[-A(x)]$$

$$\Rightarrow F_X(x) = 1 - M_{\Lambda}[-A(x)]$$

Ex 5.7 p. (66)

Let Λ have a gamma distribution and let $X|\Lambda$ have a Weibull distribution with conditional survival function $S_{X|\Lambda}(x|\lambda) = e^{-\lambda x^\gamma}$. Determine the unconditional or marginal distⁿ of X .

Ans: let $\Lambda \sim \text{gamma}(\theta, \alpha)$, $X|\Lambda \sim \text{Weibull}(\lambda, \gamma)$ See p. (469)

$$\therefore S_{X|\Lambda}(x|\lambda) = e^{-\lambda x^\gamma}$$

$$\therefore A(x) = x^\gamma$$

$$\therefore S_X(x) = E[e^{-\Lambda A(x)}]$$

$$= M_\Lambda[-A(x)]$$

$$\therefore S_X(x) = M_\Lambda(-x^\gamma)$$

and $\therefore M_\Lambda(z) = (1 - \theta z)^{-\alpha}$

$$\therefore S_X(x) = (1 + \theta x^\gamma)^{-\alpha}$$

which is a Burr distⁿ with parameters

$$\theta \rightarrow \theta^{-1/\gamma}, \quad \alpha \rightarrow \alpha \text{ (not change)} \quad \text{See p. (469)}$$

Note that when $\gamma=1$, this is an exponential mixture which is a pareto distⁿ, with parameters $\theta \rightarrow 1/\theta$, $\alpha \rightarrow \alpha$ See p. (465)

Revise
EX 1, Lec 6
For $X \sim \text{gamma}(\theta, \alpha)$
 $M_X(z) = E(e^{zX})$
 $= \frac{1}{(1-\theta z)^\alpha}, z < \frac{1}{\theta}$

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• Sec 5.2.6 Splicing "spliced distribution"

Def'n A k -component spliced distribution has a density function that can be expressed as follows: See p. (66)

$$f_X(x) = \begin{cases} a_1 f_1(x), & c_0 < x < c_1 \\ a_2 f_2(x), & c_1 < x < c_2 \\ \vdots & \vdots \\ a_k f_k(x), & c_{k-1} < x < c_k \end{cases}$$

For $j = 1, \dots, k$, each $a_j > 0$ and each $f_j(x)$ must be legitimate density function with all probability on the interval (c_{j-1}, c_j) .
Also, $a_1 + \dots + a_k = 1$.

Ex 5.8 p. (67)

Demonstrate that Model 5 in Section 2.2 is a two-component spliced model.

Ans:

For Model 5, the density f_X is

Revise Lecture 3

$$f_X(x) = \begin{cases} 0.01, & 0 \leq x < 50 \\ 0.02, & 50 \leq x < 75 \end{cases}$$

$$\Rightarrow f_X(x) = \begin{cases} 0.5(0.02), & 0 \leq x < 50 \\ 0.5(0.04), & 50 \leq x < 75 \end{cases} = \begin{cases} a_1 f_1(x), & 0 \leq x < 50 \\ a_2 f_2(x), & 50 \leq x < 75 \end{cases}$$

* It is a spliced model created by letting $f_1(x) = 0.02, 0 \leq x < 50$ which is a uniform distn on the interval from 0 to 50, and $f_2(x) = 0.04, 50 \leq x < 75$, which is a uniform distn on the interval from 50 to 75. The coefficients are $a_1 = 0.5$ and $a_2 = 0.5$.

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