King Saud University
College of Sciences
Department of Mathematics

Mid Term Exam, S3 1444
M 380 - Stochastic Processes
Time: 2 hours - Marks: 30

## Answer the following questions.

## Q1 [5]

Given the joint probability mass function of two random variables $X$ and $Y$
as in the following table:

| $X$ | -1 | 0 | 1 |
| :--- | :---: | :---: | :---: |
| -1 | $1 / 9$ | 0 | $2 / 9$ |
| 0 | $2 / 9$ | $1 / 9$ | 0 |
| 1 | 0 | $2 / 9$ | $1 / 9$ |

(i) Find $\boldsymbol{\rho}(X, Y)$
(ii) Determine whether $X$ and $Y$ are two independent random variables or not, Justify your answer.

## Q2: [4+4]

(a) Suppose that $\boldsymbol{\xi}_{1}, \boldsymbol{\xi}_{2}, \ldots$ are independent and identically distributed with $\operatorname{Pr}\left\{\boldsymbol{\xi}_{k}= \pm 1\right\}=\frac{1}{2}$. Let $N$ be independent of $\xi_{1}, \xi_{2}, \ldots$ and follow the geometric probability mass function $P_{N}(k)=\alpha(1-\alpha)^{k}$ for $k=0,1, \ldots$, where $0<\alpha<1$. Form the random sum $Z=\xi_{1}+\xi_{2}+\ldots+\xi_{N}$. Determine the mean and variance of $Z$.
(b) Let $S_{0}=0$, and for $n \geq 1$, let $S_{n}=\zeta_{1}+\zeta_{2}+\cdots+\zeta_{n}$ be the sum of $n$ independent random variables, each exponentially distributed with mean $E\left(\zeta_{k}\right)=1$. Show that $X_{n}=2^{n} e^{-S_{n}}, n \geq 0$, defines a martingale.

## Q3: [10:2+2+2+2+2]

(a) Consider the Markov chain whose transition probability matrix is given by

$$
\left.\mathbf{P}=\begin{array}{c||cccc||} 
\\
0 & 1 & 0 & 0 & 0 \\
1 & 0.2 & 0.4 & 0.3 & 0.1 \\
2 & 0.1 & 0.5 & 0.3 & 0.1 \\
3 & 0 & 0 & 0 & 1
\end{array} \right\rvert\,
$$

(i) Starting in state 1 , determine the probability that the Markov chain ends in state 0 .
(ii) Determine the mean time to absorption.
(iii) Sketch, the Markov chain diagram, and determine whether it's an absorbing chain or not.
(b) Let $X_{n}$ denote the quality of the nth item that produced in a certain factory with $X_{n}=0$ meaning "defective" and $X_{n}=1$ meaning "good". Suppose that $\left\{X_{n}\right\}$ be a Markov chain whose transition matrix is

$$
\left.\mathbf{P}=\begin{array}{cc}
0 & 1 \\
0 \| 0.89 & 0.11 \\
1 \| & \| .02 \\
0.98
\end{array} \right\rvert\,
$$

(i) What is the probability that the fourth item is defective given that the first item is good?
(ii) In the long run, what is the probability that an item produced by this system is good or it's defective?

## Q4: [7:2+2+3]

(a) Consider the problem of sending a binary message, 0 or 1 , through a signal channel consisting of several stages, where transmission through each stage is subject to a fixed probability of error $\alpha$. Suppose that $X_{0}=0$ is the signal that is sent and let $X_{n}$ be the signal that is received at the nth stage. Assume that $\left\{X_{n}\right\}$ is a Markov chain with transition probabilities $P_{00}=P_{11}=1-\alpha$ and $P_{01}=P_{10}=\alpha$, where $0<\alpha<1$.
(i) Determine $\operatorname{Pr}\left\{X_{0}=0, X_{1}=0, X_{2}=0\right\}$, the probability that no error occurs up to stage $n=2$.
(ii) Determine the probability that a correct signal is received at stage 2 .
(b) Consider a spare parts inventory model in which either 0 , 1 , or 2 repair parts are demanded in any period, with $\operatorname{Pr}\left\{\xi_{n}=0\right\}=0.3, \operatorname{Pr}\left\{\xi_{n}=1\right\}=0.2, \operatorname{Pr}\left\{\xi_{n}=2\right\}=0.5$ and suppose $s=0$ and $S=3$. Determine the transition probability matrix for the Markov chain $\left\{X_{n}\right\}$, where $X_{n}$ is defined to be the quantity on hand at the end of period $n$.

## The Model Answer

Q1: [5]

|  | ${ }^{-1}$ | 0 | 1 | $\mathrm{P}_{\mathrm{X}}(\mathrm{x})$ |
| :--- | :---: | :---: | :---: | :---: |
| -1 | $1 / 9$ | 0 | $2 / 9$ | $1 / 3$ |
| 0 | $2 / 9$ | $1 / 9$ | 0 | $1 / 3$ |
| 1 | 0 | $2 / 9$ | $1 / 9$ | $1 / 3$ |
| $\mathrm{P}_{\mathrm{Y}}(\mathrm{y})$ | $1 / 3$ | $1 / 3$ | $1 / 3$ | Sum $=1$ |

$$
\begin{aligned}
& E(X)=0, E\left(X^{2}\right)=\frac{2}{3}, \quad \operatorname{Var}(X)=\frac{2}{3} \\
& E(Y)=0, E\left(Y^{2}\right)=\frac{2}{3}, \quad \operatorname{Var}(Y)=\frac{2}{3} \\
& E(X Y)=0
\end{aligned}
$$

$\operatorname{Cov}(X, Y)=E(X Y)-E(X) E(Y)=0$

$$
\rho(X, Y)=\frac{\operatorname{Cov}(X, Y)}{\sigma_{X} \sigma_{Y}}
$$

$$
=\frac{0}{2 / 3}
$$

$$
=0
$$

$\Rightarrow X$ and $Y$ are not correlated
$\because$ For example, $P(X=0, Y=1)=0$, but $P(X=0) P(Y=1)=\frac{1}{9}$
$\Rightarrow P(X=0, Y=1) \neq P(X=0) P(Y=1)$
$\therefore X$ and $Y$ are not independent r.vs
Q2: $[4+4]$
(a) We have,

| $\boldsymbol{\xi}_{k}$ | -1 | 1 |
| :---: | :---: | :---: |
| $\operatorname{Pr}\left(\boldsymbol{\xi}_{k}\right)$ | $\frac{1}{2}$ | $\frac{1}{2}$ |

$\because \boldsymbol{\mu}=E\left(\boldsymbol{\xi}_{k}\right)=\sum_{k} \boldsymbol{\xi}_{k} \operatorname{Pr}\left(\boldsymbol{\xi}_{k}\right)$
$\therefore \mu=-1\left(\frac{1}{2}\right)+1\left(\frac{1}{2}\right)=0$
Also, $E\left(\boldsymbol{\xi}_{k}^{2}\right)=(-1)^{2}\left(\frac{1}{2}\right)+(1)^{2}\left(\frac{1}{2}\right)=1$
and $\sigma^{2}=\operatorname{Var}\left(\xi_{k}\right)=E\left(\xi_{k}^{2}\right)-\left[E\left(\xi_{k}\right)\right]^{2}=1-0=1$
$\because N \sim \operatorname{geom}(\alpha), P_{N}(k)=\alpha(1-\alpha)^{k}, k=0,1, \ldots$
$\therefore v=E(N)=\frac{1-\alpha}{\alpha}, \tau^{2}=\operatorname{Var}(N)=\frac{1-\alpha}{\alpha^{2}}$
$\because Z=\xi_{1}+\xi_{2}+\ldots+\xi_{N}$
$\therefore E(Z)=\mu v=0\left(\frac{1-\alpha}{\alpha}\right)=0$
and

$$
\begin{aligned}
\operatorname{Var}(Z) & =v \sigma^{2}+\mu^{2} \tau^{2} \\
& =\frac{1-\alpha}{\alpha}(1)+0\left(\frac{1-\alpha}{\alpha^{2}}\right)
\end{aligned}
$$

$\therefore \operatorname{Var}(Z)=\frac{1-\alpha}{\alpha}$
(b)
(1) $E\left[\left|X_{n}\right|\right]=E\left[X_{n}\right]=E\left[2^{n} e^{-S_{n}}\right]$
$=2^{n} E\left[e^{-\zeta_{1}} \ldots e^{-\zeta_{n}}\right]$
$=2^{n} E\left[e^{-\zeta_{1}}\right] \ldots E\left[e^{-\zeta_{n}}\right]$, as $\zeta_{i I_{s}}$ are independent
$=2^{n} \frac{1}{2} \cdots \cdots \frac{1}{2}=\frac{2^{n}}{2^{n}}=1$, as
$E\left[e^{-\zeta_{n}}\right]=\int_{0}^{\infty} e^{-x} e^{-x} d x$
$=\int_{0}^{\infty} e^{-2 x} d x=\frac{1}{2}$

So, $E\left[\left|X_{n}\right|\right]=1<\infty$.
(2) $E\left[X_{n+1} \mid X_{0}, \ldots . ., X_{n}\right]=E\left[2^{n+1} e^{\left.-S_{n+1} \mid X_{0}, \ldots . ., X_{n}\right], S_{n+1}=S_{n}+\zeta_{n+1}}\right.$

$$
\begin{aligned}
& =E\left[2^{n} e^{-S_{n}} 2 e^{\left.-\zeta_{n+1} \mid X_{0}, \ldots ., X_{n}\right]}\right. \\
& =2^{n} e^{-S_{n}} E\left[2 e^{\left.-\zeta_{n+1} \mid X_{0}, \ldots ., X_{n}\right]}\right. \\
& =2^{n} e^{-S_{n}} 2 E\left[e^{-\zeta_{n+1}}\right],
\end{aligned}
$$

as $\zeta_{n+1}$ is independent of $X_{i / s}$,

$$
\begin{aligned}
E\left[X_{n+1} \mid X_{0}, \ldots \ldots, X_{n}\right] & =2^{n} e^{-s_{n}} 2 \cdot \frac{1}{2} \\
& =2^{n} e^{-s_{n}} \\
& =X_{n} .
\end{aligned}
$$

We have proved from (1) and (2) that $X_{n}$ is a martingale.
Q3: $[10: 2+2+2+2+2]$
(a)

$$
\left.\mathbf{P}=\begin{array}{c||cccc||} 
\\
0 \\
1 & 1 & 0 & 0 & 0 \\
0 & 0.2 & 0.4 & 0.3 & 0.1 \\
2 \\
3 & 0.1 & 0.5 & 0.3 & 0.1 \\
0 & 0 & 0 & 1
\end{array} \right\rvert\,
$$

$u_{i}=\operatorname{pr}\left\{X_{T}=0 \mid X_{0}=i\right\}$ for $\mathrm{i}=1,2$,
and $v_{i}=\mathrm{E}\left[T \mid X_{0}=i\right] \quad$ for $\mathrm{i}=1,2$.
(i)
$u_{1}=p_{10}+p_{11} u_{1}+p_{12} u_{2}$
$u_{2}=p_{20}+p_{21} u_{1}+p_{22} u_{2}$
$\Rightarrow$

$$
\begin{aligned}
& u_{1}=0.2+0.4 u_{1}+0.3 u_{2} \\
& u_{2}=0.1+0.5 u_{1}+0.3 u_{2} \\
& \Rightarrow
\end{aligned}
$$

$$
\begin{equation*}
6 u_{1}-3 u_{2}=2 \tag{1}
\end{equation*}
$$

$5 u_{1}-7 u_{2}=-1$

Solving (1) and (2), we get

$$
u_{1}=\frac{17}{27} \text { and } u_{2}=\frac{16}{27}
$$

Starting in state 1, the probability that the Markov chain ends in state 0 is
$u_{1}=u_{10}=\frac{17}{27}=0.6296$
(ii) Also, the mean time to absorption can be found as follows

$$
\begin{aligned}
& v_{1}=1+p_{11} v_{1}+p_{12} v_{2} \\
& v_{2}=1+p_{21} v_{1}+p_{22} v_{2} \\
& \Rightarrow \\
& v_{1}=1+0.4 v_{1}+0.3 v_{2} \\
& v_{2}=1+0.5 v_{1}+0.3 v_{2} \\
& \Rightarrow
\end{aligned}
$$

$$
\begin{equation*}
6 v_{1}-3 v_{2}=10 \tag{1}
\end{equation*}
$$

$5 v_{1}-7 v_{2}=-10$
Solving (1) and (2), we get

$$
\begin{aligned}
v_{1}=v_{10} & =\frac{100}{27} \\
& \approx 3.7037
\end{aligned}
$$

(iii)


It's an absorbing Markov Chain.
(b)
(i)

$$
\begin{aligned}
& P^{2}=\left[\begin{array}{ll}
0.89 & 0.11 \\
0.02 & 0.98
\end{array}\right]\left[\begin{array}{ll}
0.89 & 0.11 \\
0.02 & 0.98
\end{array}\right] \\
& P^{2}=\left[\begin{array}{ll}
0.7943 & 0.2057 \\
0.0374 & 0.9626
\end{array}\right] \\
& P^{3}=\left[\begin{array}{ll}
0.7110 & 0.2890 \\
0.0525 & 0.9475
\end{array}\right] \\
& \operatorname{pr}\left\{X_{3}=0 \mid X_{0}=1\right\}=p_{10}^{3}=0.0525 \\
& =5.25 \%,
\end{aligned}
$$

which is the probability that the fourth item is defective given that the first item is good.
(ii) In the long run, the probability that an item produced by this system is defective is given by:

$$
\begin{aligned}
b /(a+b) & =\frac{0.02}{0.02+0.11} \\
& =\frac{2}{13}=15.38 \%
\end{aligned}
$$

In the long run, the probability that an item produced by this system is good is given by:

$$
\begin{aligned}
a /(a+b) & =\frac{0.11}{0.02+0.11} \\
& =\frac{11}{13}=84.62 \%,
\end{aligned}
$$

where $\lim _{n \rightarrow \infty} \mathrm{P}^{n}=\left[\begin{array}{cc}\frac{b}{a+b} & \frac{a}{a+b} \\ \frac{b}{a+b} & \frac{a}{a+b}\end{array}\right], \quad \mathrm{P}=\left[\begin{array}{cc}1-a & a \\ b & 1-b\end{array}\right]$
Clearly, $1-\frac{2}{13}=\frac{11}{13}=84.62 \%$
Q4: $[7: 2+2+3]$

## (a)

The transition probability matrix can be written as

$$
\mathbf{P}=\begin{array}{cc}
0 & 1 \\
0 \| 1-\alpha & \alpha \\
1 \| \begin{array}{c}
\alpha \\
\alpha
\end{array} & 1-\alpha \|
\end{array}
$$

(i) The probability that no error occurs up to stage $n=2$ is given as follows.

$$
\begin{aligned}
\operatorname{Pr}\left\{X_{0}=0, X_{1}=0, X_{2}=0\right\} & =p_{0} P_{00} P_{00}, \mathrm{p}_{0}=\operatorname{pr}\left\{X_{0}=0\right\}=1 \\
& =1 \times(1-\alpha) \times(1-\alpha) \\
& =(1-\alpha)^{2}
\end{aligned}
$$

(ii) The probability that a correct signal is received at stage 2 is given as follows.
$\operatorname{Pr}\left\{X_{0}=0, X_{1}=0, X_{2}=0\right\}+\operatorname{Pr}\left\{X_{0}=0, X_{1}=1, X_{2}=0\right\}$
$=p_{0} P_{00} P_{00}+p_{0} P_{01} P_{10}$
$=(1-\alpha)^{2}+\alpha^{2}$
$=1-2 \alpha+2 \alpha^{2}$
(b)

| -1 | 0 | 1 | 2 | 3 |
| ---: | :--- | :--- | :--- | :--- |
| -1 | 0 | 0 | 0.5 | 0.2 |
| 0 | 0.3 |  |  |  |
| 0 | 0 | 0.5 | 0.2 | 0.3 |
| 1 | 0.5 | 0.2 | 0.3 | 0 |
| 2 | 0 |  |  |  |
| 2 | 0.2 | 0.2 | 0.3 | 0 |
| 3 | 0 | 0.5 | 0.2 | 0.3 |$|$

$P_{i j}=\operatorname{Pr}\left(\xi_{n+1}=S-j\right) \quad, i \leq s \quad$ for replenishment
$P_{-1,-1}=\operatorname{Pr}\left(\xi_{n+1}=4\right)=0 \quad, P_{01}=\operatorname{Pr}\left(\xi_{n+1}=2\right)=0.5$
$P_{i j}=\operatorname{Pr}\left(\xi_{n+1}=i-j\right) \quad, s<i \leq S$ for non-replenishment
$P_{1,-1}=\operatorname{Pr}\left(\xi_{n+1}=2\right)=0.5 \quad, P_{11}=\operatorname{Pr}\left(\xi_{n+1}=0\right)=0.3, P_{21}=\operatorname{Pr}\left(\xi_{n+1}=1\right)=0.2$

