Fingdom of Saudi Cralia
Fing Saud Uniwersity
Faculty of Sciences
Final Exam-aCJU 464
1438-1 ${ }^{\text {st }}$ Semester
Duration: 3 hours


## INSTRUCTIONS

1- Please check that your exam contains 12 pages total (including the first page!!)
9 questions and a Bonus question.
2- Answer all questions.
3- No books, No notes and no phones are allowed.
4- A standard no programmable calculator is allowed.
5- Also included in this paper exam:

- A table for most used distributions.
- A Z-table.
- Some formulas are included.

| Question | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Score |  |  |  |  |  |  |  |  |  |  |

## Total score

Question 1. (3 marks)
Suppose $Z$ is the $q$-mixture of $X$ and $Y$, with $X \sim \operatorname{Negative~}-\operatorname{Binomial}(r, p)$ and $Y \sim \operatorname{Uniform}(a, b)$. Compute the mean and the mgf of $Z$.

Question 2. (2 marks)
Let consider the random variable $X$ having the following cumulative distribution function:

$$
F(x)=\left\{\begin{array}{ccc}
0 & \text { if } & x<-1 \\
(x+1) / 4 & \text { if } & -1 \leq x<1 \\
1 & \text { if } & x \geq 1
\end{array}\right.
$$

Compute the mean of $X$.

Question 3. (2+2+2=6 marks)
Suppose $X_{1}, X_{2}, \ldots X_{n}$ are iid with common distribution $X \sim \operatorname{Gamma}(\alpha .2)$.
a) Determine the distribution of $S=X_{1}+\cdots+X_{n}$.
b) Compute the mean and variance of $S$.
c) Using the normal approximation, compute $P(S<205)$ for $\alpha=4$ and $n=100$.

Question 4. (3+2=5 marks)
A total claim amounts has a $\chi^{2}(20)$ distribution. Calculate a) The mean, the variance and the skewness of $S$.
b) $P(S \geq 30)$, using the NP approximation.

Question 5. (3+2+3=8 marks)
Let $S$ be a compound random variable with claim $X$ and claim number $N$. Suppose further that $f_{X}(1)=0.5, f_{X}(2)=0.1, f_{X}(3)=0.4$ and $N \sim \operatorname{Poisson}(4)$.
a) Compute the mean, the variance and the mgf of $S$.
b) Compute, using Sparse algorithm, the mass function of $S$ at $s=0,1,2,3$.
c) Conclude the stop-loss premium $\pi(3)=E(S-3)_{+}$.

Question 6. (4 marks)
Lets be an aggregate loss random variable with a discrete frequency distribution $N$ (number of claims) defined by the table below.

| $n$ | 0 | 1 | 2 |
| :--- | :---: | :---: | :---: |
| $P(N=n)$ | 0.3 | 0.5 | 0.2 |

The severity claim $X$ has a geometric distribution with parameter $p=0.2$. Find the mass function of $S$ using the convolution formula.

Question 7. (2+2=4 marks)
The frequency distribution of an aggregate loss $S$ follows a binomial distribution with parameters $n=4$ and $p=0.3$. Loss amount has the mass distribution function: $f_{X}(0)=0.2, f_{X}(1)=0.7$ and $f_{X}(2)=0.1$. Using the Panjer's recursive formula for the binomial distribution with $a=-3 / 7$ and $b=12 / 7$,
a) Compute the mass function $f$ of $S$ at $s=1,2$.
b) Compute the mass function $f$ of $S$ for all $s \geq 3$.

Question 8. (3 marks)
In some ruin process, the individual claims have a $\operatorname{Gamma}(2,1)$ distribution. Determine the adjustment coefficient $R$ for a loading factor $\theta=1$.

Question 9. (2+3=5 marks)
Let $\left(S_{t}, t \geq 0\right)$ be a compound Poisson process with parameter $\lambda$ and exponentially distributed claims with parameter $\beta$.
Compute $E\left(S_{s} S_{t}\right)$ and $E\left(e^{2 S_{t}} \mid S_{s}\right)$ for $s<t$.

Bonus Question. (3 marks)
The frequency distribution of an aggregate loss $S$ follows a negative-binomial distribution with parameters $r=4$ and $p=0.2$. Loss amount has the mass distribution function: $f(0)=0.2 ; f(1)=0.7 ; f(2)=0.1$. Calculate the mass function of $S$ at $s=1,2$.

- NP approximation formulas.

Let a random variable $S$ with mean $\mu$, finite variance $\sigma^{2}$ and skewness $\gamma$. For $x \geq 1$ we have
$P\left(\frac{S-\mu}{\sigma} \leq x\right) \approx \phi\left(\sqrt{\frac{9}{\gamma^{2}}+\frac{6 x}{\gamma}+1}-\frac{3}{\gamma}\right)$
and for $s \geq 1$ we have
$P\left(\frac{s-\mu}{\sigma} \leq s+\frac{\gamma}{6}\left(s^{2}-1\right)\right) \approx \phi(s)$,
where $\phi$ is the standard normal cdf.

## - Translated Gamma approximation formula.

The translated gamma approximation can then be formulated as follows:

$$
\begin{align*}
& F_{S}(s) \approx G\left(s-x_{0} ; \alpha, \beta\right) \\
& \text { where } G(x ; \alpha, \beta)=\frac{1}{\Gamma(\alpha)} \int_{0}^{x} y^{\alpha-1} \beta^{\alpha} \mathrm{e}^{-\beta y} \mathrm{~d} y, \quad x \geq 0 \tag{2.56}
\end{align*}
$$

Here $G(x ; \alpha, \beta)$ is the gamma cdf. We choose $\alpha, \beta$ and $x_{0}$ such that the first three moments are the same, hence $\mu=x_{0}+\frac{\alpha}{\beta}, \sigma^{2}=\frac{\alpha}{\beta^{2}}$ and $\gamma=\frac{2}{\sqrt{\alpha}}$ (see Table A), so

$$
\begin{equation*}
\alpha=\frac{4}{\gamma^{2}}, \quad \beta=\frac{2}{\gamma \sigma} \quad \text { and } \quad x_{0}=\mu-\frac{2 \sigma}{\gamma} \tag{2.57}
\end{equation*}
$$

- Panjer's recursive formula.


## Theorem 3.5.1 (Panjer's recursion)

Consider a compound distribution with integer-valued non-negative claims with pdf $p(x), x=0,1,2, \ldots$, for which, for some real $a$ and $b$, the probability $q_{n}$ of having $n$ claims satisfies the following recursion relation

$$
\begin{equation*}
q_{n}=\left(a+\frac{b}{n}\right) q_{n-1}, \quad n=1,2, \ldots \tag{3.26}
\end{equation*}
$$

Then the following relations for the probability of a total claim equal to $s$ hold:

$$
\begin{align*}
f(0) & = \begin{cases}\operatorname{Pr}[N=0] & \text { if } \quad p(0)=0 \\
\mathrm{~m}_{N}(\log p(0)) & \text { if } \quad p(0)>0\end{cases}  \tag{3.27}\\
f(s) & =\frac{1}{1-a p(0)} \sum_{h=1}^{s}\left(a+\frac{b h}{s}\right) p(h) f(s-h), \quad s=1,2, \ldots
\end{align*}
$$

Table A The most frequently used discrete and continuous distributions

| Distribution | Density \& support | Moments \& cumulants | Mgf |
| :---: | :---: | :---: | :---: |
| $\begin{aligned} & \operatorname{Binomial}(n, p) \\ & (0<p<1, n \in \mathbb{N}) \end{aligned}$ | $\begin{gathered} \binom{n}{x} p^{x}(1-p)^{n-x} \\ x=0,1, \ldots, n \end{gathered}$ | $\begin{aligned} & \mathrm{E}=n p, \operatorname{Var}=n p(1-p), \\ & \gamma=\frac{n p(1-p)(1-2 p)}{\sigma^{3}} \end{aligned}$ | $\left(1-p+p \mathrm{e}^{t}\right)^{n}$ |
| Bernoulli $(p)$ | $\equiv \operatorname{Binomial}(1, p)$ |  |  |
| Poisson( $\lambda$ ) $(\lambda>0)$ | $\mathrm{e}^{-\lambda} \frac{\lambda^{x}}{x!}, x=0,1, \ldots$ | $\begin{aligned} & \mathrm{E}=\mathrm{Var}=\lambda \\ & \gamma=1 / \sqrt{\lambda} \\ & \kappa_{j}=\lambda, j=1,2, \ldots \end{aligned}$ | $\exp \left[\lambda\left(e^{t}-1\right)\right]$ |
| $\begin{aligned} & \text { Negative } \\ & \quad \text { binomial }(r, p) \\ & (r>0,0<p<1) \end{aligned}$ | $\begin{gathered} \binom{r+x-1}{x} p^{r}(1-p)^{x} \\ x=0,1,2, \ldots \end{gathered}$ | $\begin{aligned} & \mathrm{E}=r(1-p) / p \\ & \mathrm{Var}=\mathrm{E} / p, \\ & \gamma=\frac{(2-p)}{p \sigma} \end{aligned}$ | $\left(\frac{p}{1-(1-p) \mathrm{e}^{t}}\right)^{r}$ |
| Geometric( $p$ ) | $\equiv$ Negative binomial $(1, p)$ |  |  |
| $\begin{aligned} & \operatorname{Uniform}(a, b) \\ & (a<b) \end{aligned}$ | $\frac{1}{b-a} ; a<x<b$ | $\begin{aligned} & \mathrm{E}=(a+b) / 2 \\ & \mathrm{Var}=(b-a)^{2} / 12, \\ & \gamma=0 \end{aligned}$ | $\frac{\mathrm{e}^{b t}-\mathrm{e}^{a t}}{(b-a) t}$ |
| $\begin{aligned} & \mathrm{N}\left(\mu, \sigma^{2}\right) \\ & (\sigma>0) \end{aligned}$ | $\frac{1}{\sigma \sqrt{2 \pi}} \exp \frac{-(x-\mu)^{2}}{2 \sigma^{2}}$ | $\begin{aligned} & \mathrm{E}=\mu, \mathrm{Var}=\sigma^{2}, \gamma=0 \\ & \left(\kappa_{j}=0, j \geq 3\right) \end{aligned}$ | $\exp \left(\mu t+\frac{1}{2} \sigma^{2} t^{2}\right)$ |
| $\begin{aligned} & \operatorname{Gamma}(\alpha, \beta) \\ & (\alpha, \beta>0) \end{aligned}$ | $\frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha-1} \mathrm{e}^{-\beta x}, x>0$ | $\begin{aligned} & \mathrm{E}=\alpha / \beta, \operatorname{Var}=\alpha / \beta^{2}, \\ & \gamma=2 / \sqrt{\alpha} \end{aligned}$ | $\left(\frac{\beta}{\beta-t}\right)^{\alpha}(t<\beta)$ |
| Exponential $(\beta)$ | $\equiv \operatorname{gamma}(1, \beta)$ |  |  |
| $\chi^{2}(k)(k \in \mathbb{N})$ | $\equiv \operatorname{gamma}(k / 2,1 / 2)$ |  |  |
| $\begin{aligned} & \text { Inverse } \\ & \quad \text { Gaussian }(\alpha, \beta) \\ & (\alpha>0, \beta>0) \end{aligned}$ | $\begin{gathered} \frac{\alpha x^{-3 / 2}}{\sqrt{2 \pi \beta}} \exp \left(\frac{-(\alpha-\beta x)^{2}}{2 \beta x}\right) \\ F(x)=\Phi\left(\frac{-\alpha}{\sqrt{\beta x}}+\sqrt{\beta x}\right) \end{gathered}$ | $\begin{aligned} & \mathrm{E}=\alpha / \beta, \text { Var }=\alpha / \beta^{2}, \\ & \gamma=3 / \sqrt{\alpha} \\ & +\mathrm{e}^{2 \alpha} \Phi\left(\frac{-\alpha}{\sqrt{\beta x}}-\sqrt{\beta x}\right), \end{aligned}$ | $\begin{aligned} & \mathrm{e}^{\alpha(1-\sqrt{1-2 t / \beta})} \\ & (t \leq \beta / 2) \\ & \quad x>0 \end{aligned}$ |

## Standard Normal Probabilities



Table entry for $z$ is the area under the standard normal curve to the left of $z$.

| $z$ | .00 | .01 | .02 | .03 | .04 | .05 | .06 | .07 | .08 | .09 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0.0 | .5000 | .5040 | .5080 | .5120 | .5160 | .5199 | .5239 | .5279 | .5319 | .5359 |
| 0.1 | .5398 | .5438 | .5478 | .5517 | .5557 | .5596 | .5636 | .5675 | .5714 | .5753 |
| 0.2 | .5793 | .5832 | .5871 | .5910 | .5948 | .5987 | .6026 | .6064 | .6103 | .6141 |
| 0.3 | .6179 | .6217 | .6255 | .6293 | .6331 | .6368 | .6406 | .6443 | .6480 | .6517 |
| 0.4 | .6554 | .6591 | .6628 | .6664 | .6700 | .6736 | .6772 | .6808 | .6844 | .6879 |
| 0.5 | .6915 | .6950 | .6985 | .7019 | .7054 | .7088 | .7123 | .7157 | .7190 | .7224 |
| 0.6 | .7257 | .7291 | .7324 | .7357 | .7389 | .7422 | .7454 | .7486 | .7517 | .7549 |
| 0.7 | .7588 | .7611 | .7642 | .7673 | .704 | .7734 | .7764 | .7794 | .7823 | .7852 |
| 0.8 | .7881 | .7910 | .7939 | .7967 | .7995 | .8023 | .8051 | .8078 | .8106 | .8133 |
| 0.9 | .8159 | .8186 | .8212 | .8238 | .8264 | .8289 | .8315 | .8340 | .8365 | .8389 |
| 1.0 | .8413 | .8438 | .8461 | .8485 | .8508 | .8531 | .8554 | .8577 | .8599 | .8621 |
| 1.1 | .8643 | .8665 | .8686 | .8708 | .8729 | .8749 | .8770 | .8790 | .8810 | .88030 |
| 1.2 | .8849 | .8869 | .8888 | .8907 | .8925 | .8944 | .8962 | .8980 | .8997 | .9015 |
| 1.3 | .9032 | .9049 | .9066 | .9082 | .9099 | .9115 | .9131 | .9147 | .9162 | .9177 |
| 1.4 | .9192 | .9207 | .9222 | .9236 | .9251 | .9265 | .9279 | .9292 | .9306 | .9319 |
| 1.5 | .9332 | .9345 | .9357 | .9370 | .9382 | .9394 | .9406 | .9418 | .9429 | .9441 |
| 1.6 | .9452 | .9463 | .9474 | .9484 | .9495 | .9505 | .9515 | .9525 | .9535 | .9545 |
| 1.7 | .9554 | .9564 | .9573 | .9582 | .9591 | .9599 | .9608 | .9616 | .9625 | .9633 |
| 1.8 | .9641 | .9649 | .9656 | .9664 | .9671 | .9678 | .9686 | .9693 | .9699 | .9706 |
| 1.9 | .9713 | .9719 | .9726 | .9732 | .9738 | .9744 | .9750 | .9756 | .9761 | .9767 |
| 2.0 | .9772 | .9778 | .9783 | .9788 | .9793 | .9798 | .9803 | .9808 | .9812 | .9817 |
| 2.1 | .9821 | .9826 | .9830 | .9834 | .9838 | .9842 | .9846 | .9850 | .9854 | .9857 |
| 2.2 | .9861 | .9864 | .9868 | .9871 | .9875 | .9878 | .9881 | .9884 | .9887 | .9890 |
| 2.3 | .9893 | .9896 | .9898 | .9901 | .9904 | .9906 | .9909 | .9911 | .9913 | .9916 |
| 2.4 | .9918 | .9920 | .9922 | .9925 | .9927 | .9929 | .9931 | .9932 | .9934 | .9936 |
| 2.5 | .9938 | .9940 | .9941 | .9943 | .9945 | .9946 | .9948 | .9949 | .9951 | .9952 |
| 2.6 | .9953 | .9955 | .9956 | .9957 | .9959 | .9960 | .9961 | .9962 | .9963 | .9964 |
| 2.7 | .9965 | .9966 | .9967 | .9968 | .9969 | .9970 | .9971 | .9972 | .9973 | .9974 |
| 2.8 | .9974 | .9975 | .9976 | .9977 | .9977 | .9978 | .9979 | .9979 | .9980 | .9981 |
| 2.9 | .9981 | .9982 | .9982 | .9983 | .9984 | .9984 | .9985 | .9985 | .9986 | .9986 |
| 3.0 | .9987 | .9987 | .9987 | .9988 | .9988 | .9989 | .9989 | .9989 | .9990 | .9990 |
| 3.1 | .9990 | .9991 | .9991 | .9991 | .9992 | .9992 | .9992 | .9992 | .9993 | .9993 |
| 3.2 | .9993 | .9993 | .9994 | .9994 | .9994 | .9994 | .9994 | .9995 | .9995 | .9995 |
| 3.3 | .9995 | .9995 | .9995 | .9996 | .9996 | .9996 | .9996 | .9996 | .9996 | .9997 |
| 3.4 | .9997 | .9997 | .9997 | .9997 | .9997 | .9997 | .9997 | .9997 | .9997 | .9998 |

