



Answer the following questions:

Q1: [6+3]

a) Determine the distribution function, mean and variance corresponding to the

triangular density $f(x) = \begin{cases} x & \text{for } 0 \leq x \leq 1 \\ 2-x & \text{for } 1 \leq x \leq 2 \\ 0 & \text{elsewhere} \end{cases}$

b) Find the moment generating function $M_X(t)$ of X , where $X \sim \text{Uniform}(a,b)$

Q2: [4+4]

a) Determine the mean and the median of an exponentially distributed random variable with parameter λ

b) If $X \sim \text{Bin}(p, N)$ and $N \sim \text{Poisson}(\lambda)$. What is the marginal distribution for X ?

Q3: [4+4]

a) Let X and Y are jointly distributed random variables having the density

function $f_{XY}(x,y) = \frac{1}{y} e^{-(x/y)-y}$ for $x,y > 0$ find $f_{XY}(x|y)$

b) Suppose that a number of miles that a car can run before its battery wears out is exponentially distributed with an average value of 10000 miles. If a person desires to take a 5000-mile trip, what is the probability that he will be able to complete his trip without having to replace the car battery?

Q1

9) $F(x) = \int_0^x t dt = \frac{x^2}{2}, 0 \leq x < 1$

$F(x) = \int_0^x (2-t) dt = [2t - \frac{t^2}{2}]_0^x$
 $= \frac{1}{2}(x^2 - 4x + 3)$

$= \frac{1}{2}[1 - (x-2)^2]$
 $1 \leq x < 2$

$\therefore F(x) = \begin{cases} 0 & x < 0 \\ \frac{1}{2}x^2 & 0 \leq x < 1 \\ \frac{1}{2}[1 - (x-2)^2] & 1 \leq x < 2 \\ 1 & x \geq 2 \end{cases}$

$E(X) = \int_{-\infty}^{\infty} x f(x) dx$
 $= \int_0^1 x(x) dx + \int_1^2 x(2-x) dx$
 $= [\frac{x^3}{3}]_0^1 + [\frac{2x^2}{2} - \frac{x^3}{3}]_1^2$

$E(X) = 1$

$Var(X) = E(X^2) - \mu^2$

$E(X^2) = \int x^2 f(x) dx$
 $= \int_0^1 x^3 dx + \int_1^2 x^2(2-x) dx$

$E(X^2) = [\frac{x^4}{4}]_0^1 + [\frac{2x^3}{3} - \frac{x^4}{4}]_1^2$
 $= 7/6$

$V(X) = \frac{1}{2} + \frac{14}{3} - 1 - 1$
 $= \frac{1}{6}$

b) $M_X(t) = E(e^{xt})$

$= \int_{-\infty}^{\infty} e^{xt} f(x) dx$

$= \int_a^b e^{xt} \frac{1}{b-a} dx$
 $X \sim \text{Uniform}(a,b)$

$= \frac{1}{b-a} \left[\frac{e^{xt}}{t} \right]_a^b$

$= \frac{1}{b-a} \left(\frac{e^{bt} - e^{at}}{t} \right)$

$\therefore M_X(t) = \frac{1}{t(b-a)} (e^{bt} - e^{at}), t \neq 0$

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(8)

Q(2)

a) $X \sim \text{exp}(\lambda)$, $\lambda > 0$

$$f(x|\lambda) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{else} \end{cases}$$

$$\mu = E(X) = \lambda \int_0^{\infty} x e^{-\lambda x} dx$$

$$\text{let } u = \lambda x \Rightarrow x = \frac{u}{\lambda}, dx = \frac{du}{\lambda}$$

$$\mu = \frac{\lambda}{\lambda^2} \int_0^{\infty} u e^{-u} du$$

$$\therefore \mu = \frac{1}{\lambda} \Gamma(2) = \frac{1}{\lambda} \text{ (mean)}$$

For $\text{pr}(X \leq a) \geq \frac{1}{2}$

$$1 - e^{-\lambda a} \geq \frac{1}{2}$$

$$e^{-\lambda a} \leq \frac{1}{2}$$

$$e^{\lambda a} \geq 2$$

$$\therefore a \geq \frac{\ln 2}{\lambda} \quad (1)$$

For $\text{pr}(X \geq a) \geq \frac{1}{2}$

$$e^{-\lambda a} \geq \frac{1}{2}$$

$$\therefore a \leq \frac{\ln 2}{\lambda} \quad (2)$$

$$(1), (2) \Rightarrow \text{Median} = \frac{\ln 2}{\lambda}$$

i.e. Median < Mean for the r.v. $X \sim \text{exp}(\lambda)$

b) $X \sim \text{Bin}(p, N)$, $N \sim \text{poisson}(\lambda)$

$$P(X=N) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$P_N(n) = \frac{e^{-\lambda} \lambda^n}{n!}, n=0,1,2,\dots$$

$$\text{Pr}(X=x) = \sum_{n=0}^{\infty} P(x|n) P_N(n)$$

total prob.

Subs. (1), (2) in (3) \Rightarrow

$$\text{Pr}(X=x) = \sum_{n=x}^{\infty} \frac{n!}{x!(n-x)!} p^x (1-p)^{n-x}$$

$$= \frac{\lambda^x e^{-\lambda} p^x}{x!} \sum_{n=x}^{\infty} \frac{[\lambda(1-p)]^{n-x}}{(n-x)!}$$

$$= \frac{(\lambda p)^x e^{-\lambda}}{x!} \sum_{r=0}^{\infty} \frac{[\lambda(1-p)]^r}{r!}$$

where $r = n - x$

$$\text{Pr}(X=x) = \frac{(\lambda p)^x e^{-\lambda} e^{\lambda(1-p)}}{x!}$$

$$\text{Pr}(X=x) = \frac{(\lambda p)^x e^{-\lambda p}}{x!}, x=0,1,2,\dots$$

$\therefore X \sim \text{poisson}(\lambda p)$ with mean λp #

Q3

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(4)

$$a) f_{X|Y}(x|y) = \frac{1}{y} e^{-(x/y)-y}$$

for $x, y > 0$

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

$$f_Y(y) = \int_0^{\infty} f(x,y) dx$$

$$f_Y(y) = \int_0^{\infty} \frac{1}{y} e^{-(x/y)-y} dx$$

$$f_Y(y) = \frac{e^{-y}}{y} \int_0^{\infty} e^{-x/y} dx$$

$$f_Y(y) = \frac{e^{-y}}{y} \left[\frac{e^{-x/y}}{-1/y} \right]_0^{\infty}$$

$$f_Y(y) = e^{-y} [0 + 1] = e^{-y}, y > 0$$

Note: $\lim_{x \rightarrow \infty} e^{-x} = 0$

b) $X \sim \exp\left(\frac{1}{10000}\right)$

$$Pr(X > 5000)$$

$$= e^{-\frac{5000}{10000}}$$

$$= e^{-0.5} \approx 0.6065$$

(4)

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