Chapter 10: One- and Two-Sample Tests of Hypotheses:

Consider a population with some unknown parameter θ . We are interested in testing (confirming or denying) some conjectures about θ . For example, we might be interested in testing the conjecture that $\theta = \theta_0$, where θ_0 is a given value.

10.1-10.3: Introduction +:

- A statistical hypothesis is a conjecture concerning (or a statement about) the population.
- For example, if θ is an unknown parameter of the population, we may be interested in testing the conjecture that $\theta > \theta_o$ for some specific value θ_o .
- We usually test the null hypothesis:

 $H_0: \theta = \theta_0$ (Null Hypothesis) Against one of the following alternative hypotheses:

 $H_{1}: \begin{cases} \theta \neq \theta_{o} \\ \theta > \theta_{o} \\ \theta < \theta_{o} \end{cases}$ (Alternative Hypothesis or Research Hypothesis)

• Possible situations in testing a statistical hypothesis:

| | H _o is true | H _o is false |
|--------------------------|------------------------|-------------------------|
| Accepting H _o | Correct Decision | Type II error |
| | | (β) |
| Rejecting H _o | Type I error | Correct Decision |
| | (α) | |

Type I error = Rejecting H_o when H_o is true Type II error = Accepting H_o when H_o is false P(Type I error) = P(Rejecting $H_o | H_o$ is true) = α P(Type II error) = P(Accepting $H_o | H_o$ is false) = β

• The level of significance of the test: $\alpha = P(Type \ I \ error) = P(Rejecting \ H_o \ | \ H_o \ is \ true)$ • One-sided alternative hypothesis:

 $H_o: \theta = \theta_o \text{ or } H_o: \theta = \theta_o$

 $H_1\!\!:\theta > \theta_o \qquad \qquad H_1\!\!:\theta < \theta_o$

- Two-sided alternative hypothesis:
 - $H_{o}: \theta = \theta_{o}$ $H_{1}: \theta \neq \theta_{o}$
- The test procedure for rejecting H_o (accepting H_1) or accepting H_o (rejecting H_1) involves the following steps:
- 1. Determining a test statistic (T.S.)
- 2. Determining the significance level α $\alpha = 0.01, 0.025, 0.05, \text{ or } 0.10$
- 3. Determining the rejection region (R.R.) and the acceptance region (A.R.) of H_0 .

R.R. of H_{o} depends on H_{1} and α

- H_1 determines the direction of the R.R. of H_o
- α determines the size of the R.R. of H_o



4. Decision:

We reject H_o (accept H_1) if the value of the T.S. falls in the R.R. of H_o , and vice versa.

<u>10.5: Single Sample: Tests Concerning a Single Mean</u> (Variance Known):

Suppose that $X_1, X_2, ..., X_n$ is a random sample of size *n* from distribution with mean μ and (known) variance σ^2 . Recall:

•
$$E(\overline{X}) = \mu_{\overline{X}} = \mu$$

•
$$Var(\overline{X}) = \sigma_{\overline{X}}^2 = \frac{\sigma^2}{n}$$

•
$$\overline{X} \sim N\left(\mu, \frac{\sigma}{\sqrt{n}}\right) \Leftrightarrow Z = \frac{\overline{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$$

• Let μ_0 be a given known value.

• Test Procedure:

| Hypotheses | $H_o: \mu = \mu_o$ | $H_o: \mu = \mu_o$ | $H_o: \mu = \mu_o$ | | |
|-------------------|---|---|--|--|--|
| | $H_1: \mu \neq \mu_o$ | $H_1: \mu > \mu_o$ | $H_1: \mu < \mu_o$ | | |
| Test Statistic | , | $\overline{X} - \mu_0$ N(0.1) | | | |
| (T.S.) | | $L = \frac{1}{\sigma/\sqrt{n}} \sim N(0,1)$ | | | |
| R.R. and A.R. | \square | \frown | \square | | |
| of H _o | $1-\alpha$ | | | | |
| | $\alpha/2$ A R of Ha | $1-\alpha$ | α $1-\alpha$ | | |
| | R.R. $Z_{1-\alpha/2}$ $Z_{\alpha/2}$ R.R. | $A.R. of H_0$ Z_{α} $R.R.$ | R.R. $Z_{1-\alpha}$ A.R. of H _o | | |
| | $= - Z_{\alpha/2}$ | • — a of H _o | $=-Z_{\alpha}$ | | |
| Decision: | Reject H_0 (and accept H_1) at the significance level α if: | | | | |
| | $Z > Z_{\alpha/2}$ | $Z > Z_{\alpha}$ | $Z < -Z_{\alpha}$ | | |
| | or $Z < -Z_{\alpha/2}$ | | | | |
| | Two-Sided Test | One-Sided Test | One-Sided Test | | |

Example 10.3:

A random sample of 100 recorded deaths in the United States during the past year showed an average of 71.8 years. Assuming a population standard deviation of 8.9 year, does this seem to indicate that the mean life span today is greater than 70 years? Use a 0.05 level of significance.

Solution:

 $n=100, \quad \overline{X} = 71.8, \quad \sigma=8.9$ $\mu=$ average (mean) life span $\mu_0=70$

Department of Statistics and O.R. - 102 -

King Saud University

Hypotheses:

H_o:
$$\mu = 70$$

H₁: $\mu > 70$
T.S. :
 $Z = \frac{\overline{X} - \mu_o}{\sigma / \sqrt{n}} = \frac{71.8 - 70}{8.9 / \sqrt{100}} = 2.02$
Level of significance:
 $\alpha = 0.05$
R R ·



К.К.:

$$Z_{\alpha} = Z_{0.05} = 1.645$$

 $Z > Z_{\alpha} = Z_{0.05} = 1.645$

Decision:

Since Z=2.02 \in R.R., i.e., Z=2.02>Z_{0.05}, we reject H₀ at α =0.05 and accept H₁: μ > 70. Therefore, we conclude that the mean life span today is greater than 70 years.

Example 10.4:

A manufacturer of sports equipment has developed a new synthetic fishing line that he claims has a mean breaking strength of 8 kilograms with a standard deviation of 0.5 kilograms. Test the hypothesis that $\mu=8$ kg against the alternative that $\mu \neq 8$ kg if a random sample of 50 lines is tested and found to have a mean breaking strength of 7.8 kg. Use a 0.01 level of significance.

Solution:



Department of Statistics and O.R. - 103 -King Saud University Decision:

Since $Z=-2.83 \in R.R.$, i.e., $Z=-2.83 < -Z_{0.005}$, we reject H_o at $\alpha=0.01$ and accept H_1 : $\mu \neq 8$. Therefore, we conclude that the claim is not correct.

<u>10.7: Single Sample: Tests on a Single Mean (Variance Unknown):</u>

Suppose that $X_1, X_2, ..., X_n$ is a random sample of size *n* from normal distribution with mean μ and unknown variance σ^2 . Recall:

•
$$T = \frac{\overline{X} - \mu}{S / \sqrt{n}} \sim t(v)$$
; $v = n - 1$
• $S = \sqrt{\sum_{i=1}^{n} (X_i - \overline{X})^2 / (n - 1)}$

• Test Procedure:

| Hypotheses | $H_o: \mu = \mu_o$ | $H_o: \mu = \mu_o$ | $H_o: \mu = \mu_o$ |
|------------------------|--|--|--|
| | $H_1: \mu \neq \mu_o$ | $H_1: \mu > \mu_o$ | $H_1: \mu < \mu_o$ |
| Test Statistic | 7 | $\overline{X} - \mu_0$ at $(n-1)$ | |
| (T.S.) | 1 | $-\frac{1}{S/\sqrt{n}} \sim t(n-1)$ | |
| R.R. and | \land | \square | \land |
| A.R. of H _o | | | |
| | $\alpha/2$ $1-\alpha$ $\alpha/2$ | $1-\alpha$ α | α $1-\alpha$ |
| | A.R. of H _o | | |
| | $ \begin{array}{ccc} \text{R.R.} & t_{1-\alpha/2} & t_{\alpha/2} & \text{of } H_{0} \\ \hline &t & \end{array} $ | A.R. of H_0 \mathbf{t}_{α} of H_0 | $\int_{\alpha} \int_{\alpha} \int_{\alpha$ |
| | - τ _{α/2} | | |
| Decision: | Reject H_0 (and accept H_1) at the significance level α if: | | |
| | $T > t_{\alpha/2}$ | $T > t_{\alpha}$ | $T < -t_{\alpha}$ |
| | or $T < -t_{\alpha/2}$ | | |
| | Two-Sided Test | One-Sided Test | One-Sided Test |

Example 10.5:

... If a random sample of 12 homes included in a planned study indicates that vacuum cleaners expend an average of 42 kilowatt-hours per year with a standard deviation of 11.9 kilowatt-hours, does this suggest at the 0.05 level of significance that the vacuum cleaners expend, on the average, less than 46

Department of Statistics and O.R. – 104 – King Saud University

kilowatt-hours annually? Assume the population of kilowatt-hours to be normal.

Solution:

 $.n=12, \overline{X}=42, S=11.9, \alpha=0.05$

µ=average (mean) kilowatt-hours annual expense of a vacuum
cleaner

µ₀=46

Hypotheses:

H_o: $\mu = 46$ H₁: $\mu < 46$ T.S. : $T = \frac{\overline{X} - \mu_0}{S / \sqrt{n}} = \frac{42 - 46}{11.9 / \sqrt{12}} = -1.16$; $\nu = df = n - 1 = 11$ $- t_{\alpha} = -t_{0.05} = -1.796$



Decision:

Since T= $-1.16 \notin R.R.$ (T= $-1.16 \in A.R.$), we do not reject H_o at α =0.05 (i.e, accept H_o: μ = 46) and reject H₁: μ < 46. Therefore, we conclude that μ is not less than 46 kilowatt-hours.

10.8: Two Samples: Tests on Two Means:

Recall: For two independent samples:

• If σ_1^2 and σ_2^2 are known, then we have:

$$Z = \frac{(\overline{X}_1 - \overline{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0, 1)$$

• If σ_1^2 and σ_2^2 are unknown but $\sigma_1^2 = \sigma_2^2 = \sigma^2$, then we have:

$$T = \frac{(\overline{X}_1 - \overline{X}_2) - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t(n_1 + n_2 - 2)$$

Where the pooled estimate of
$$\sigma^2$$
 is

$$S_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}$$

The degrees of freedom of S_p^2 is $v=n_1+n_2-2$. Now, suppose we need to test the null hypothesis

Department of Statistics and O.R. – 105 – King Saud University

 $H_o: \mu_1 = \mu_2 \quad \Leftrightarrow \quad H_o: \mu_1 - \mu_2 = 0$ Generally, suppose we need to test

H_o: $\mu_1 - \mu_2 = d$ (for some specific value *d*) Against one of the following alternative hypothesis

H₁:
$$\begin{cases} \mu_1 - \mu_2 \neq d \\ \mu_1 - \mu_2 > d \\ \mu_1 - \mu_2 < d \end{cases}$$

| Hypotheses | $H_0: \mu_1 - \mu_2 = d$ | $H_0: \mu_1 - \mu_2 = d$ | $H_0: \mu_1 - \mu_2 = d$ | |
|------------------------------------|---|---|--|--|
| | $\mathbf{H}_1: \boldsymbol{\mu}_1 - \boldsymbol{\mu}_2 \neq d$ | H ₁ : $\mu_1 - \mu_2 > d$ | H ₁ : $\mu_1 - \mu_2 < d$ | |
| Test Statistic (T.S.) | $Z = \frac{(\overline{X}_1 - \overline{X}_2) - d}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0, 1) \qquad \text{{if } } \sigma_1^2 \text{ and } \sigma_2^2 \text{ are known}$ | | | |
| | or $T = \frac{(\overline{X}_1 - \overline{X}_2) - d}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t(n)$ | (n_1+n_2-2) {if $\sigma_1^2 = \sigma_2^2$ | $\frac{2}{2} = \sigma^2$ is unknown} | |
| R.R. and A.R. of H _o | $\alpha/2 \qquad 1 - \alpha \qquad \alpha/2 \\ A.R. \text{ of } H_0 \qquad \alpha/2 \\ A.R. \text{ of } H_0 \qquad \alpha/2 \\ = - Z_{\alpha/2} \qquad Z_{\alpha/2} \qquad R.R. \\ \sigma H_0 \qquad = - Z_{\alpha/2} \\ Or \qquad $ | Or $A.R. of H_0$ $Z_{\alpha} R.R. of H_0$ $C_{\alpha} R.R. of H_0$ $C_{\alpha} R.R. of H_0$ | $\alpha \qquad 1-\alpha$ $R.R. Z_{1-\alpha} \qquad A.R. \text{ of } H_0$ $= -Z_{\alpha}$ Or $\alpha \qquad 1-\alpha$ $R.R. t_{1-\alpha} \qquad A.R. \text{ of } H_0$ $= -t_{\alpha}$ | |
| Decision: | Reject H _o (and accept H ₁) at the significance level α if: | | | |
| | T.S. ∈ R.R. | T.S. ∈ R.R. | T.S. ∈ R.R. | |
| | Two-Sided Test | One-Sided Test | One-Sided Test | |

Example10.6:

An experiment was performed to compare the abrasive wear of

Department of Statistics and O.R. – 106 – King Saud University

two different laminated materials. Twelve pieces of material 1 were tested by exposing each piece to a machine measuring wear. Ten pieces of material 2 were similarly tested. In each case, the depth of wear was observed. The samples of material 1 gave an average wear of 85 units with a sample standard deviation of 4, while the samples of materials 2 gave an average wear of 81 and a sample standard deviation of 5. Can we conclude at the 0.05 level of significance that the mean abrasive wear of material 1 exceeds that of material 2 by more than 2 units? Assume populations to be approximately normal with equal variances.

Solution:

| | Material 1 | material 2 | |
|---|---|-------------------------------------|------------------------------------|
| | $n_1 = 12$ | $n_2 = 10$ | |
| | $\overline{X}_1 = 85$ | $\overline{X}_2 = 81$ | |
| _ | S ₁ =4 | S ₂ =5 | |
| Hypotheses: | | \sim | |
| $H_0: \mu_1 = \mu_2$ | + 2 (<i>d</i> =2) | | \backslash |
| $H_1: \mu_1 > \mu_2$ | + 2 | | \backslash |
| Or equivalently | у, | 0.95 | 0.05 |
| $H_0: \mu_1 - \mu_2 =$ | = 2 (<i>d</i> =2) | | |
| $H_1: \mu_1 - \mu_2$ | > 2 | | to or R.R. |
| Calculation: | | | *0.05 of H _o = 1 725 |
| α=0.05 | | | - 1.720 |
| $s^2 = (n_1 - 1)S_1^2$ | $+(n_2-1)S_2^2$ (12-1 | $(4)^{2} + (10 - 1)(5)^{2}$ | - 20.05 |
| $S_p = \frac{n_1 + 1}{n_1 + 1}$ | $n_2 - 2$ – – – – – – – – – – – – – – – – – – | 12 + 10 - 2 | 20.03 |
| $S_{p}=4.478$ | | | |
| $v = n_1 + n_2 - 2 = 12$ | 2+10-2=20 | | |
| $t_{0.05} = 1.725$ | | | |
| T.S.: | | | |
| $T - \frac{(\overline{X}_1 - \overline{X}_2)}{(\overline{X}_1 - \overline{X}_2)}$ | $(\overline{K}_2) - d$ (85 – 81) |)-2 -1.04 | |
| $I = \frac{S_p \sqrt{\frac{1}{n_p}}}{S_p \sqrt{\frac{1}{n_p}}}$ | $\frac{1}{1} - \frac{1}{n_2} = \frac{1}{(4.478)\sqrt{\frac{1}{1}}}$ | $\frac{1}{2} + \frac{1}{10} = 1.04$ | |
| Decision: | | | |

Since T=1.04 \in A.R. (T=1.04< $t_{0.05} = 1.725$), we accept (do not reject) H_o and reject H₁: $\mu_1 - \mu_2 > 2$ at $\alpha = 0.05$.

10.11 One Sample: Tests on a Single Proportion:



Recall:

• .*p* = Population proportion of elements of Type *A* in the population

• =
$$\frac{A}{A+B} = \frac{no. \ of \ elements \ of \ type \ A \ in \ the \ population}{Total \ no. \ of \ elements}$$

- .*n* = sample size
- X =no. of elements of type A in the sample of size n.
- \hat{p} = Sample proportion elements of Type *A* in the sample = $\frac{X}{2}$

• For large *n*, we have

$$Z = \frac{\hat{p} - p}{\sqrt{\frac{pq}{n}}} = \frac{X - np}{\sqrt{npq}} \sim N(0, 1) \quad \text{(Approximately, } q = 1 - p\text{)}$$

| Hypotheses | $H_o: p = p_o$ | $H_o: p = p_o$ | $H_o: p = p_o$ | |
|--|--|--|--|--|
| | $H_1: p \neq p_o$ | $H_1: p > p_0$ | $H_1: p < p_0$ | |
| Test Statistic (T.S.) | $Z = \frac{\hat{p} - p_o}{\sqrt{\frac{p_o q_o}{n}}} =$ | $=\frac{X-np_o}{\sqrt{np_oq_o}} \sim N(0,1) ($ | <i>q</i> _o =1– <i>p</i> _o) | |
| R.R. and A.R. of H _o | $\begin{array}{c} \alpha/2 & 1-\alpha \\ A.R. \text{ of } H_0 \\ \hline R.R. \\ of H_0 \\ = -Z_{\alpha/2} \end{array} \begin{array}{c} \alpha/2 \\ Z_{\alpha/2} \\ R.R. \\ of H_0 \\ R.R. \\ of H_0 \end{array}$ | $1 - \alpha$ A.R. of H _o $Z_{\alpha} R.R. of H_{o}$ | $\begin{array}{c} \alpha & 1-\alpha \\ R.R. & Z_{1-\alpha} \\ of H_0 \\ = -Z_{\alpha} \end{array}$ | |
| Decision: | Reject H_0 (and accept H_1) at the significance level α if: | | | |
| | $Z > Z_{\alpha/2}$ or $Z < -Z_{\alpha/2}$ | $\overline{Z > Z_{\alpha}}$ | $Z < -Z_{\alpha}$ | |
| | Two-Sided Test | One-Sided Test | One-Sided Test | |
| Department of Statistics and O.R. – 108 – King Saud University | | | | |

Example 10.10:

A builder claims that heat pumps are installed in 70% of all homes being constructed today in the city of Richmond. Would you agree with this claim if a random survey of new homes in the city shows that 8 out of 15 homes had heat pumps installed? Use a 0.10 level of significance.

Solution:

p = Proportion of homes with heat pumps installed in the city.<math>n=15

X= no. of homes with heat pumps installed in the sample = 8

 \hat{p} = proportion of homes with heat pumps installed in the

sample =
$$\frac{X}{n} = \frac{8}{15} = 0.5333$$

Hypotheses:

H_o:
$$p = 0.7$$
 ($p_0=0.7$)
H₁: $p \neq 0.7$
Level of significance:

α=0.10 T.S.:

$$Z = \frac{\hat{p} - p_o}{\sqrt{\frac{p_o q_o}{n}}} = \frac{0.5333 - 0.70}{\sqrt{\frac{(0.7)(0.3)}{15}}} = -1.41$$



or

$$Z = \frac{X - np_o}{\sqrt{np_oq_o}} = \frac{8 - (15)(0.7)}{\sqrt{(15)(0.7)(0.3)}} = -1.41$$
$$Z_{\alpha/2} = Z_{0.05} = 1.645$$

Decision:

Since $Z = -1.41 \in A.R.$, we accept (do not reject) H_0 : p=0.7 and reject H_1 : $p \neq 0.7$ at $\alpha=0.1$. Therefore, we agree with the claim.

Example 10.11: Reading Assignment

| Department of Statistics and O.R. | - 109 - | King Saud University |
|-----------------------------------|---------|----------------------|

10.12 Two Samples: Tests on Two Proportions:



Suppose that we have two populations:

- $.p_1$ = proportion of the 1-st population.
- $.p_2$ = proportion of the 2-nd population.
- We are interested in comparing *p*₁ and *p*₂, or equivalently, making inferences about *p*₁−*p*₂.
- We <u>independently</u> select a random sample of size n_1 from the 1-st population and another random sample of size n_2 from the 2-nd population:
- Let $X_1 = no$. of elements of type *A* in the 1-st sample.
- Let $X_2 = no.$ of elements of type A in the 2-nd sample.
- $\hat{p}_1 = \frac{X_1}{n_1}$ = proportion of the 1-st sample
- $\hat{p}_2 = \frac{X_2}{n_2}$ = proportion of the 2-nd sample
- The sampling distribution of $\hat{p}_1 \hat{p}_2$ is used to make inferences about $p_1 p_2$.
- For large n_1 and n_2 , we have

$$Z = \frac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\sqrt{\frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}}} \sim N(0, 1) \quad \text{(Approximately)}$$

Department of Statistics and O.R. – 110 – King Saud University

Suppose we need to test:

$$\mathbf{H}_{0}: p_{1} = p_{2}$$
$$\mathbf{H}_{1}: \begin{cases} p_{1} \neq p_{2} \\ p_{1} > p_{2} \\ p_{1} < p_{2} \\ p_{1} < p_{2} \end{cases}$$

Or, equivalently,

$$H_{0}: p_{1} - p_{2} = 0$$

$$H_{1}: \begin{cases} p_{1} - p_{2} \neq 0 \\ p_{1} - p_{2} > 0 \\ p_{1} - p_{2} < 0 \end{cases}$$

Note, under H_0 : $p_1 = p_2 = p$, the pooled estimate of the proportion p is:

$$\hat{p} = \frac{X_1 + X_2}{n_1 + n_2}$$
 ($\hat{q} = 1 - \hat{p}$)

The test statistic (T.S.) is

$$Z = \frac{(\hat{p}_{1} - \hat{p}_{2})}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{n_{1}} + \frac{1}{n_{2}}\right)}} \sim N(0,1)$$
Hypotheses $H_{0}: p_{1} - p_{2} = 0$ $H_{0}: p_{1} - p_{2} = 0$ $H_{1}: p_{1} - p_{2} \neq 0$ $H_{1}: p_{1} - p_{2} > 0$ $H_{1}: p_{1} - p_{2} < 0$
Test Statistic (T.S.)
$$Z = \frac{(\hat{p}_{1} - \hat{p}_{2})}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{n_{1}} + \frac{1}{n_{2}}\right)}} \sim N(0,1)$$
R.R. and A.R. of H_{0} u^{2} $A_{R. of H_{0}}$ u^{2} u^{2

Department of Statistics and O.R. – 111 – King Saud University

| $Z > Z_{\alpha/2}$ | $Z > Z_{\alpha}$ | $Z < -Z_{\alpha}$ |
|------------------------|-----------------------|-----------------------|
| or $Z < -Z_{\alpha/2}$ | | |
| Two-Sided Test | One-Sided Test | One-Sided Test |

Example 10.12:

A vote is to be taken among the residents of a town and the surrounding county to determine whether a proposed chemical plant should be constructed. The construction site is within the town limits and for this reason many voters in the county feel that the proposal will pass because of the large proportion of town voters who favor the construction. To determine if there is a significant difference in the proposal, a poll is taken. If 120 of 200 town voters favor the proposal and 240 of 500 county voters favor it, would you agree that the proportion of town voters favoring the proposal and 240 of 500 county voters favoring the proposal is higher than the proportion of county voters? Use a 0.025 level of significance.

Solution:

 $p_1 =$ proportion of town voters favoring the proposal $p_2 =$ proportion of county voters favoring the proposal $\hat{p}_1 =$ sample proportion of town voters favoring the proposal $\hat{p}_2 =$ sample proportion of county voters favoring the proposal

Town
 $n_1 = 200$ County
 $n_2 = 500$ $X_1 = 120$ $X_2 = 240$ $\hat{p}_1 = \frac{X_1}{n_1} = \frac{120}{200} = 0.60$ $\hat{p}_2 = \frac{X_2}{n_2} = \frac{240}{500} = 0.48$ $\hat{q}_1 = 1 - 0.60 = 0.40$ $\hat{q}_2 = 1 - 0.48 = 0.52$

The pooled estimate of the proportion p is:

 $\hat{p} = \frac{X_1 + X_2}{n_1 + n_2} = \frac{120 + 240}{200 + 500} = 0.51$ $\hat{q} = 1 - 0.51 = 0.49$ Hypotheses: H₀: $p_1 = p_2$

Department of Statistics and O.R. - 112 -

H₁:
$$p_1 > p_2$$

or
H₀: $p_1 - p_2 = 0$
H₁: $p_1 - p_2 > 0$
Level of significance:
 $\alpha = 0.025$
 $Z_{\alpha} = Z_{0.025} = 1.96$
T.S.:
 $Z = \frac{(\hat{p}_1 - \hat{p}_2)}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} = \frac{(0.60 - 0.48)}{\sqrt{(0.51)(0.49)\left(\frac{1}{200} + \frac{1}{500}\right)}} = 2.869$

Decision:

Since Z= 2.869 \in R.R. (Z= 2.869> Z_{α}= Z_{0.025}= 1.96), we reject H_o: $p_1 = p_2$ and accept H₁: $p_1 > p_2$ at α =0.025. Therefore, we agree that the proportion of town voters favoring the proposal is higher than the proportion of county voters.