

Solution of the First Midterm exam: Stochastic Processes: Math. 380 (25%)

Sunday, March 19, 2017 (1 – 3) PM

Problem 1. (5 marks)

Let the random variable  $X$  have the p.m.f.  $f(k) = \frac{1}{3}$ ,  $k \in S_X = \{-1, 0, 1\}$ . Let  $Y = X^2$ ,

1. (1 mark) Find the distribution of  $Y$ ,

**Solution:**  $S_Y = \{0, 1\}$ , and the p.m.f. is given by

$$f(0) = P(Y = 0) = P(X = 0) = \frac{1}{3}$$

and

$$\begin{aligned} f(1) &= P(Y = 1) = P(X^2 = 1) = P(X = -1 \text{ or } X = 1) \\ &= P(X = -1) + P(X = 1) = \frac{1}{3} + \frac{1}{3} = \frac{2}{3}. \end{aligned}$$

2. (1 mark) Calculate the expectation of  $X$  and  $Y$ .

**Solution:**

$$E[X] = \sum_{k=-1}^1 k \frac{1}{3} = \frac{1}{3}(-1 + 0 + 1) = 0 \quad \text{and} \quad E[Y] = \sum_{k=0}^1 k f(k) = \frac{1}{3} \times 0 + \frac{2}{3} \times 1 = \frac{2}{3}$$

3. (1 mark) Deduce the variance of  $X$ .

**Solution:**

$$\text{Var}(X) = E[X^2] - (E[X])^2 = E[Y] = \frac{2}{3}.$$

4. (1 mark) Find the covariance of  $X$  and  $Y$ .

**Solution:**

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = E[X^3] = \frac{1}{3}(-1 + 0 + 1) = 0.$$

5. (1 mark) Are  $X$  and  $Y$  independent? Explain your answer.

**Solution:**  $X$  and  $Y$  are dependent because  $Y = X^2$ .

Problem 2. (5 marks)

1. Let the joint p.m.f. of  $X$  and  $Y$  be defined by  $f(i, j) = \frac{i+j}{32}$ ,  $i = 1, 2$ , and  $j = 1, 2, 3, 4$ .

- (a) (1 mark) Find  $f_X(i)$  the marginal p.m.f. of  $X$ .

$$f_X(i) = \sum_{j=1}^4 f(i, j) = \sum_{j=1}^4 \frac{i+j}{32} = \frac{5+2i}{16}$$

(b) (1 mark) Find  $f_Y(j)$  the marginal p.m.f. of  $Y$ .

$$f_Y(j) = \sum_{i=1}^2 f(i, j) = \sum_{i=1}^2 \frac{i+j}{32} = \frac{3+2j}{32}$$

(c) (1 mark) Find  $P(Y = 2X)$ . Let us first remark that

$$\begin{aligned} \{Y = 2X\} &= \{Y = 2X\} \cap (\{X = 1\} \cup \{X = 2\}) \\ &= (\{Y = 2X\} \cap \{X = 1\}) \cup (\{Y = 2X\} \cap \{X = 2\}) \end{aligned}$$

hence

$$\begin{aligned} P(Y = 2X) &= P(\{Y = 2X\} \cap \{X = 1\}) + P(\{Y = 2X\} \cap \{X = 2\}) \\ &= P(Y = 2; X = 1) + P(Y = 4; X = 2) \\ &= \frac{1+2}{32} + \frac{2+4}{32} = \frac{9}{32}. \end{aligned}$$

(d) (1 mark) Find  $P(X \leq 3 - Y)$ .

$$\begin{aligned} P(X \leq 3 - Y) &= P(X \leq 3 - Y; X = 1) + P(X \leq 3 - Y; X = 2) \\ &= P(1 \leq 3 - Y; X = 1) + P(2 \leq 3 - Y; X = 2) \\ &= P(Y \leq 2; X = 1) + P(Y \leq 1; X = 2) \\ &= P(Y = 1; X = 1) + P(Y = 2; X = 1) + P(Y = 1; X = 2) \\ &= \frac{1+1}{32} + \frac{1+2}{32} + \frac{2+1}{32} = \frac{8}{32}. \end{aligned}$$

2. There are **eight** similar chips in a bowl: **three** marked  $(0, 0)$ , **two** marked  $(1, 0)$ , two marked  $(0, 1)$ , and one marked  $(1, 1)$ . A player selects a chip at random and is given the **sum** of the two coordinates in dollars. If  $X$  and  $Y$  represent those two coordinates, respectively, their joint p.m.f. is  $f(i, j) = \frac{3-i-j}{8}$ ,  $i = 0, 1$  and  $j = 0, 1$ .

(a) (1 mark) Calculate the expected given sum in dollars.

$$E[X + Y] = \sum_{i=0}^1 \sum_{j=0}^1 (i+j) \frac{3-i-j}{8} = \frac{3}{4}$$

### Problem 3. (5 marks)

A **discrete** is a random variable whose range is either finite or countably infinite. A **continuous** random variable is a random variable whose range is an interval in  $\mathbb{R}$ : A **mixed random variable** is partially discrete and partially continuous.

State whether the random variables are discrete, continuous or mixed.

1. (1 mark) A coin is tossed ten times. The random variable  $X$  is the number of tails that are noted. Explain your answer.

**Solution:** The r.v.  $X$  follows a binomial with parameter 10 and  $\frac{1}{2}$  since it has two alternative outcomes and can take the values  $\{0, 1, \dots, 10\}$ .

2. (1 mark) A light bulb is burned until it burns out. The random variable  $Y$  is its lifetime in hours. Explain your answer.

**Solution:** The life time a light bulb is a continuous random variable since it can take any positive value.

3. (1 mark)  $Z : ]0; 1[ \rightarrow \mathbb{R}$  where

$$Z(s) = \begin{cases} 1 - s & \text{if } 0 < s < \frac{1}{2} \\ \frac{1}{2} & \text{if } \frac{1}{2} \leq s < 1 \end{cases} \quad \text{Explain your answer.}$$

**Solution:** The r.v.  $Z$  is a mixed distribution since it takes a continuous value between  $0 < s \leq \frac{1}{2}$  and takes one single value on the interval  $\frac{1}{2} \leq s < 1$ .

4. (1 mark) Find the expectation of  $X$  and  $Z$ ;

**Solution:** We know that  $X \hookrightarrow \mathcal{B}(10, \frac{1}{2})$  the,  $E[X] = 10 \cdot \frac{1}{2} = 5$ . and

$$E[Z] = \int_0^1 z(s) ds = \int_0^{\frac{1}{2}} (1 - s) ds + \int_{\frac{1}{2}}^1 \frac{1}{2} ds = \frac{5}{8}.$$

5. (1 mark) Find the variance of  $X$  and  $Z$  **canceled**.

#### Problem 4. (6 marks)

Let  $X$  and  $Y$  be two random variables with the joint density

$$f(x, y) = \begin{cases} a(e^{-2x-3y}) & \text{if } 0 < x \text{ and } 0 < y \\ 0 & \text{elsewhere} \end{cases}$$

1. (1 mark) Find the value of  $a$ .

**Solution:** We have

$$\int_0^{\infty} \int_0^{\infty} a(e^{-2x-3y}) dx dy = 1 \iff a = 6$$

2. (1 mark) Calculate the marginal densities of  $X$  and  $Y$ .

**Solution:** We know

$$f_X(x) = \int f(x, y) dy = \int_0^{\infty} 6(e^{-2x-3y}) dy = 2e^{-2x}$$

and

$$f_Y(y) = \int f(x, y) dx = \int_0^{\infty} 6(e^{-2x-3y}) dx = 3e^{-3y}.$$

3. (1 mark) Are  $X$  and  $Y$  independent?

**Solution:** We have  $f_X(x)f_Y(y) = f(x, y) = f_{(X,Y)}(x, y)$  then  $X$  and  $Y$  are independent.

4. (1 mark) Find  $Cov(X, Y)$ .

**Solution:** By definition

$$Cov(X, Y) = E[XY] - E[X]E[Y] = 0 \text{ since } X \text{ and } Y \text{ are independent.}$$

5. (1 mark) Set  $Z = X + Y$ , Calculate the c.d.f.  $F_Z(z)$  of  $Z$ .

**Solution:** By definition

$$\begin{aligned} F_Z(z) &= P(Z \leq z) = P(X + Y \leq z) = \int \int_{\{x+y \leq z\}} f(x, y) dx dy \\ &= \int_0^z \left( \int_0^{z-x} f(x, y) dy \right) dx = \int_0^z 2e^{-2x} \left( \int_0^{z-x} 3e^{-3y} dy \right) dx \\ &= 1 + 2e^{-3z} - 3e^{-2z}. \end{aligned}$$

6. (1 mark) Find the p.d.f.  $f_Z(z)$  of  $Z$ .

**Solution:** We know that  $f_Z(z) = F'_Z(z) = 6(e^{-2z} - e^{-3z})$ .

**Problem 5. (6 mark)**

1. An insurance policy is written to cover a loss,  $X$ , where  $X$  has a uniform distribution on  $[0, 1000]$ . The policy has a deductible,  $d$ , and the expected payment under the policy is 25% of what it would be with no deductible.

(a) (2 mark) Calculate  $d$ .

**Solution:** Let  $Y$  be the payment under the deductible policy and  $X$  be the payment under policy with no deductible. So

$$E[Y] = E[\max(X - d, 0)] = \frac{25}{100}E[X] \quad (1)$$

but we know that

$$E[\max(X - d, 0)] = \int_d^{1000} (x - d) \frac{1}{1000} dx = \frac{1}{2000} (d - 1000)^2$$

and

$$E[X] = \int_0^{1000} x \frac{1}{1000} dx = 500.$$

Now substituting in the equation (??) we get

$$\begin{aligned} \frac{1}{2000} (d - 1000)^2 &= \frac{25}{100} \times 500 = 125 \\ \iff (d - 1000)^2 &= 250000 \\ \iff |d - 1000| &= \sqrt{250000} = 500 \end{aligned}$$

Then the possible solution are  $d = 500$  and  $d = 1500$ . But since  $d$  should be less than the maximum value of  $X$  we finally get as the unique solution  $d = 500$ .

2. A group insurance policy covers the medical claims of the employees of a small company. The value,  $V$ , of the claims made in one year is described by  $V = 100,000Y$  where  $Y$  is a random variable with density function

$$f(y) = \begin{cases} c(1 - y)^4 & \text{if } 0 < y < 1 \\ 0 & \text{elsewhere} \end{cases}$$

where  $c$  is a constant.

(a) (2 mark) Calculate the conditional probability that  $V$  exceeds 40,000, given that  $V$  exceeds 10,000.

**Solution:** First we have  $\int_0^1 c(1 - y)^4 dy = 1$  implies that  $\frac{1}{5}c = 1$ , then  $c = 5$ . And

$$\begin{aligned} P(V > 40,000 \mid V > 10,000) &= \frac{P(V > 40,000, V > 10,000)}{P(V > 10,000)} \\ &= \frac{P(V > 40,000, V > 10,000)}{P(V > 10,000)} \\ &= \frac{P(V > 40,000)}{P(V > 10,000)} = \frac{P(Y > \frac{4}{10})}{P(Y > \frac{1}{10})} \\ &= \frac{\int_{\frac{2}{5}}^1 5(1 - y)^4 dy}{\int_{\frac{1}{10}}^1 5(1 - y)^4 dy} = \frac{\frac{243}{3125}}{\frac{59049}{100000}} = \frac{32}{243} \simeq 0.132 \end{aligned}$$

3. The stock prices of two companies at the end of any given year are modeled with random variables  $X$  and  $Y$  that follow a distribution with joint density function

$$f(x, y) = \begin{cases} 2x, & \text{if } 0 < x < 1 \text{ and } x < y < x + 1, \\ 0, & \text{otherwise.} \end{cases}$$

Recall that

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} \quad \text{and} \quad f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)}$$

where  $f_X(x) = \int f(x, y)dy$  and  $f_Y(y) = \int f(x, y)dx$ . The conditional expectation of  $Y$  given  $X = x$  is defined by

$$E[Y | X = x] = \int y f_{Y|X}(y|x) dy = h(x) \quad \text{and} \quad E[X | Y = y] = \int x f_{X|Y}(x|y) dx = g(y)$$

and

$$\text{Var}(Y | X = x) = E[(Y - E[Y | X = x])^2 | X = x] = E[Y^2 | X = x] - (E[Y | X = x])^2.$$

- (a) **(2 mark)** Determine the conditional variance of  $Y$  given that  $X = x$ .

**Solution:** We have

$$\text{Var}(Y | X = x) = E[(Y - E[Y | X = x])^2 | X = x] = E[Y^2 | X = x] - (E[Y | X = x])^2.$$

So we need first to calculate  $f_{Y|X}(y|x)$ . By definition we have

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_X(x)} \quad \text{and} \quad f_X(x) = \int_x^{x+1} 2x dy = 2x$$

hence

$$f_{Y|X}(y|x) = \begin{cases} 1, & \text{if } 0 < x < 1 \text{ and } x < y < x + 1, \\ 0, & \text{otherwise.} \end{cases}$$

Then

$$\begin{aligned} E[Y|X = x] &= \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy = \int_x^{x+1} y dy \\ &= \left[ \frac{y^2}{2} \right]_x^{x+1} = \frac{1}{2} [(x+1)^2 - x^2] \\ &= \frac{1}{2} [2x + 1] = x + \frac{1}{2}. \end{aligned}$$

and

$$\begin{aligned} E[Y^2|X = x] &= \int_{-\infty}^{\infty} y^2 f_{Y|X}(y|x) dy = \int_x^{x+1} y^2 dy \\ &= \left[ \frac{y^3}{3} \right]_x^{x+1} = \frac{1}{3} [(x+1)^3 - x^3] = x^2 + x + \frac{1}{3} \end{aligned}$$

Therefore

$$\text{Var}(Y | X = x) = x^2 + x + \frac{1}{3} - \left(x + \frac{1}{2}\right)^2 = \frac{1}{12}.$$

**Optional : Get marks as much as you can (bonus)**

1. (1 mark) Let  $X, Y, Z$  be independent Poisson random variables with  $E[X] = 3$ ;  $E[Y] = 1$ ; and  $E[Z] = 4$ : Find  $P(X + Y + Z \leq 1)$ ?

**Solution:** We have

$$P(X + Y + Z \leq 1) = P(X + Y + Z = 0) + P(X + Y + Z = 1).$$

But we know that

$$\begin{aligned} P(X + Y + Z = 0) &= P(X = 0, Y = 0, Z = 0) \\ &= P(X = 0) P(Y = 0) P(Z = 0) \quad \text{since } X, Y, Z \text{ are independent} \\ &= e^{-3} e^{-1} e^{-4} = e^{-8} \end{aligned}$$

and

$$\begin{aligned} P(X + Y + Z = 1) &= P(X = 0, Y = 0, Z = 1) \\ &\quad + P(X = 0, Y = 1, Z = 0) \\ &\quad + P(X = 1, Y = 0, Z = 0) \\ &= P(X = 0) P(Y = 0) P(Z = 1) \\ &\quad + P(X = 0) P(Y = 1) P(Z = 0) \\ &\quad + P(X = 1) P(Y = 0) P(Z = 0) \\ &= e^{-3} e^{-1} e^{-4} 4 + e^{-3} e^{-1} e^{-4} + 3e^{-3} e^{-1} e^{-4} \\ &= 8e^{-8}. \end{aligned}$$

where we have used that fact that  $X, Y, Z$  are independent. Finally we get

$$P(X + Y + Z \leq 1) = e^{-8} + 8e^{-8} = 9e^{-8}.$$

2. Assume  $X_1, \dots, X_n$  are independent and identically distributed random variables with a common c.d.f.  $F(x)$ : Define  $U$  and  $L$  as follows  $U = \max(X_1, \dots, X_n)$  and  $L = \min(X_1, \dots, X_n)$ .

- (a) (1 mark) Write in terms of the events  $\{X_i \leq u\}$  the event  $\{U \leq u\}$  for any  $u \in \mathbb{R}$ ,

**Solution:**

- (b) (1 mark) Deduce c.d.f. of  $U$ .

**Solution:**

- (c) (1 mark) Find  $P(L > \ell)$ , for any  $\ell \in \mathbb{R}$ .

**Solution:**

- (d) (1 mark) and Deduce c.d.f. of  $L$ .

**Solution:**

- (e) (1 mark) Are  $U$  and  $L$  independent ?

**Solution:**