

Chapter 1

Q1.19)

Grade point average. The director of admissions of a small college selected 120 students at random from the new freshman class in a study to determine whether a student's grade point average (OPA) at the end of the freshman year (Y) can be predicted from the ACT test score (X). The results of the study follow. Assume that first-order regression model (1.1) is appropriate.

- Obtain the least squares estimates of β_0 and β_1 , and state the estimated regression function.
- Plot the estimated regression function and the data. "Does the estimated regression function appear to fit the data well?"
- Obtain a point estimate of the mean freshman OPA for students with ACT test score $X = 30$.
- What is the point estimate of the change in the mean response when the entrance test score increases by one point?

Solution:

$$\bar{X} = 24.725, \bar{Y} = 3.07405$$

$$\sum_{i=1}^{n=120} (X_i - \bar{X})(Y_i - \bar{Y}) = 92.40565$$

$$\sum_{i=1}^{n=120} (X_i - \bar{X})^2 = 2379.925$$

$$\sum_{i=1}^{n=120} (Y_i - \bar{Y})^2 = 49.40545$$

$$b_1 = \hat{\beta}_1 = \frac{\sum_{i=1}^{n=120} (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^{n=120} (X_i - \bar{X})^2} = \frac{92.40565}{2379.925} = 0.038827$$

$$b_0 = \hat{\beta}_0 = \bar{Y} - b_1\bar{X} = 3.07405 - 0.038827 * 24.725 = 2.114049$$

$$\hat{Y} = 2.114 + 0.0388 X$$

At $X=30$

$$\hat{Y}_h = 2.114 + 0.0388 (30) = 3.278863$$

when the entrance test score increases by one point, the mean response increase by 0.038827.

Q1.20)

Copier maintenance. The Tri-City Office Equipment Corporation sells an imported copier on a franchise basis and performs preventive maintenance and repair service on this copier. The data below have been collected from 45 recent calls on users to perform routine preventive maintenance service; for each call, X is the number of copiers serviced and Y is the total number of minutes spent by the service person. Assume that first-order regression model (1.1) is appropriate.

(مصنع يعمل على الصنعة الوقائية)

X هو عدد الناسخات الخدمات

Y هو العدد الإجمالي للدقائق التي يقضيها الشخص الخدمة

- Obtain the estimated regression function.
- Plot the estimated regression function and the data. How well does the estimated regression function fit the data?
- Interpret b_0 in your estimated regression function. Does b_0 provide any relevant information here? Explain.
- Obtain a point estimate of the mean service time when $X = 5$ copiers are serviced.

Solution:

$$\bar{X} = 5.11111, \bar{Y} = 76.26667$$

$$\sum_{i=1}^{n=120} (X_i - \bar{X})(Y_i - \bar{Y}) = 5118.667$$

$$\sum_{i=1}^{n=120} (X_i - \bar{X})^2 = 340.4444$$

$$\sum_{i=1}^{n=120} (Y_i - \bar{Y})^2 = 80376.8$$

$$b_1 = \hat{\beta}_1 = \frac{\sum_{i=1}^{n=120} (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^{n=120} (X_i - \bar{X})^2} = 15.03525$$

$$b_0 = \hat{\beta}_0 = \bar{Y} - b_1\bar{X} = -0.58016$$

$$\hat{Y} = -0.58016 + 15.03525 X$$

At X=5

$$\hat{Y}_h = -0.58016 + 15.03525 (5) = 74.59608$$

Q1.21) (H.W)

Airfreight breakage. A substance used in biological and medical research is shipped by airfreight to users in cartons of 1,000 ampules. The data below, involving 10 shipments, were collected on the number of times the carton was transferred from one aircraft to another over the shipment route (X) and the number of ampules found to be broken upon arrival (Y). Assume that first-order regression model (1.1) is appropriate.

- a. Obtain the estimated regression function. Plot the estimated regression function and the data. Does a linear regression function appear to give a good fit here?
- b. Obtain a point estimate of the expected number of broken ampules when $X = 1$ transfer is made.
- c. Estimate the increase in the expected number of ampules broken when there are 2 transfers as compared to 1 transfer.
- d. Verify that your fitted regression line goes through the point (\bar{X}, \bar{Y}) .

Q1.22)

Plastic hardness. Refer to Problems 1.3 and 1.14. Sixteen batches of the plastic were made, and from each batch one test item was molded. Each test item was randomly assigned to one of the four predetermined time levels, and the hardness was measured after the assigned elapsed time. The results are shown below; X is the elapsed time in hours? and Y is hardness in Brinell units. Assume that first-order regression model (1.1) is appropriate.

- a. Obtain the estimated regression function. Plot the estimated regression function and the data. Does a linear regression function appear to give a good fit here?
- b. Obtain a point estimate of the mean hardness when $X = 40$ hours.
- c. Obtain a point estimate of the change in mean hardness when X increases by 1 hour.

Solution:

$$\bar{X} = 28, \bar{Y} = 225.5625$$

$$\sum_{i=1}^{n=120} (X_i - \bar{X})(Y_i - \bar{Y}) = 2604$$

$$\sum_{i=1}^{n=120} (X_i - \bar{X})^2 = 1280$$

$$\sum_{i=1}^{n=120} (Y_i - \bar{Y})^2 = 5443.938$$

$$b_1 = \widehat{\beta}_1 = \frac{\sum_{i=1}^{n=120} (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^{n=120} (X_i - \bar{X})^2} = 2.034375$$

$$b_0 = \widehat{\beta}_0 = \bar{Y} - b_1 \bar{X} = 168.6$$

$$\hat{Y} = 168.6 + 2.034375 X$$

At $X=40$

$$\hat{Y}_h = 168.6 + 2.034375 (40) = 249.975$$

Q1.24) Refer to Copier maintenance Problem 1.20.

a Obtain the residuals e_i and the sum of the squared residuals $\sum e_i^2$. What is the relation between the sum of the squared residuals here and the quantity Q in (1.8)?

b. Obtain point estimates of σ^2 and σ . In what units is σ expressed?

$$\sum e_i^2 = 3416.377$$

$$\sum e_i^2 = Q$$

$$\widehat{\sigma}^2 = \frac{\sum e_i^2}{n-2} = \frac{3416.377}{43} = 79.45063 = MSE$$

$$\sigma = \sqrt{MSE} = \sqrt{79.45063}$$

Q1.25) (H.W) Refer to Airfreight breakage Problem 1.21.

- a. Obtain the residual for the first case. What is its relation to e_1 ?
- b. Compute $\sum e_i^2$ and MSE . What is estimated by MSE ?

Q1.26) (H.W) Refer to Plastic hardness Problem 1.22.

- a. Obtain the residuals e_j . Do they sum to zero in accord with (1.17)?
- b. Estimate σ^2 and σ . In what units is σ expressed?

Q1.21) Solution (H.W)

Airfreight breakage. A substance used in biological and medical research is shipped by airfreight to users in cartons of 1,000 ampules. The data below, involving 10 shipments, were collected on the number of times the carton was transferred from one aircraft to another over the shipment route (X) and the number of ampules found to be broken upon arrival (Y). Assume that first-order regression model (1.1) is appropriate.

- a. Obtain the estimated regression function. Plot the estimated regression function and the data. Does a linear regression function appear to give a good fit here?

$$\hat{Y} = 10.2 + 4.0X$$

- b. Obtain a point estimate of the expected number of broken ampules when $X = 1$ transfer is made.

If $X=1$

Then

$$\hat{Y}_h = 10.2 + 4.0(1) = 14.20$$

- c. Estimate the increase in the expected number of ampules broken when there are 2 transfers as compared to 1 transfer.

$$\hat{Y}_{h2} = 10.2 + 4.0(2) = 18.20$$

$$\hat{Y}_{h1} = 10.2 + 4.0(1) = 14.20$$

$$\hat{Y}_{h2} - \hat{Y}_{h1} = b_1 = 4.0$$

- d. Verify that your fitted regression line goes through the point (\bar{X}, \bar{Y}) .

$$\bar{X} = 1, \bar{Y} = 14.2$$

$$(\bar{X}, \bar{Y}) = (1, 14.2)$$

If $X=1$

Then

$$\hat{Y}_h = 10.2 + 4.0(1) = 14.20$$

Then we can say the regression line goes through the point $(\bar{X}, \bar{Y}) = (1, 14.2)$

Chapter 2

We assume that the normal error regression model is applicable. This model is:

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

where:

β_0 and β_1 , are parameters

X_i are known constants

ε_i are independent $N(0, \sigma^2)$

$$E(Y_i) = \beta_0 + \beta_1 X_i$$

Sampling Distribution of $\widehat{\beta}_1$

$$\widehat{\beta}_1 = b_1 = \sum_{i=1}^n \frac{(X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$E(\widehat{\beta}_1) = \beta_1$$

$$\sigma^2(\widehat{\beta}_1) = \frac{\sigma^2}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$s^2(\widehat{\beta}_1) = \frac{MSE}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$\frac{b_1 - \beta_1}{s(b_1)} \sim t_{(n-2)}$$

Confidence Interval for β_1

$$P \left[b_1 - t_{(1-\frac{\alpha}{2}, n-2)} s(b_1) \leq \beta_1 \leq b_1 + t_{(1-\alpha/2, n-2)} s(b_1) \right] = 1 - \alpha$$

C.I $(1 - \alpha)\%$ for β_1

$$b_1 - t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) \leq \beta_1 \leq b_1 + t_{(1-\alpha/2, n-2)}s(b_1)$$

Tests Concerning β_1

1. Hypothesis		
$H_0: \beta_1 = \beta_{10}$ $H_1: \beta_1 \neq \beta_{10}$	$H_0: \beta_1 = \beta_{10}$ $H_1: \beta_1 > \beta_{10}$	$H_0: \beta_1 = \beta_{10}$ $H_1: \beta_1 < \beta_{10}$
2. Test statistic		
$T_0 = \frac{b_1 - \beta_{10}}{s(b_1)}$		
3. Decision: Reject H_0 if		
$ T_0 > t_{(1-\frac{\alpha}{2}, n-2)}$	$T_0 > t_{(1-\alpha, n-2)}$	$T_0 < t_{(\alpha, n-2)}$
P-value: Reject H_0 if $p - value < \alpha$		
p-value = $2P(t_{(n-2)} > T_0)$	p - value = $P(t_{(n-2)} > T_0)$	p - value = $P(t_{(n-2)} < T_0)$

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Q2.4. Refer to **Grade point average** Problem 1.19.

a. Obtain a **99 percent confidence interval for β_1** . Interpret your confidence interval. Does it include zero? Why might the director of admissions be interested in whether the confidence interval includes zero?

Solution:

By using Minitab:

Stat → Regression → Regression → Fit Regression Mode

Minitab - Grade point Average.MPJ - [Session]

File Edit Data Calc Stat Graph Editor Tools Window Help Assistant

Regression Analysis: Yi versus Xi

Analysis of Variance

Source	DF	Seq SS	Contribution	Adj SS	Adj MS	F-Value	P-Value
Regression	1	3.588	7.26%	3.588	3.5878	9.24	0.003
Xi	1	3.588	7.26%	3.588	3.5878	9.24	0.003
Error	118	45.618	92.74%	45.618	0.3863		
Lack-of-Fit	19	6.466	12.13%	6.466	0.3414	0.86	0.632
Pure Error	99	39.332	78.61%	39.332	0.3973		
Total	119	49.406	100.00%				

Model Summary

S	R-sq	R-sq(adj)	R-sq(pred)
0.623125	7.26%	6.42%	47.6103

Coefficients

Term	Coef	SE Coef	99% CI	T-Value	P-Value	VIF
Constant	2.114	0.321	(1.274, 2.954)	6.59	0.000	
Xi	0.0388	0.0388	(0.0004, 0.0732)	1.04	0.003	1.00

Regression Equation

$$Y_i = 2.114 + 0.0388 X_i$$

Fits and Diagnostics for Unusual Observations

Obs	Yi	Fit	SE Fit	99% CI	Resid	Std Resid	Del	Resid	SI	Cook's d
2	3.855	2.658	0.145	(2.269, 3.046)	1.227	2.03	2.06	0.0566650	0.12	
9	0.500	3.240	0.079	(3.034, 3.446)	-2.740	-4.43	-4.43	0.0160124	0.16	
101	1.941	3.085	0.087	(2.936, 3.234)	-1.244	-2.00	-2.03	0.0083681	0.02	
102	1.583	2.613	0.105	(2.343, 3.083)	-1.230	-2.00	-2.03	0.0273363	0.06	
104	3.716	3.672	0.143	(3.099, 3.847)	0.043	0.60	0.60	0.0026962	0.00	
115	1.486	3.318	0.098	(3.060, 3.576)	-1.832	-2.98	-3.08	0.0248782	0.11	

Obs

Obs	DFITS	R	X
2	0.502705	R	X
9	-0.656775	R	
101	-0.186512	R	
102	-0.339912	R	
104	0.094161	X	
115	-0.492302	R	

R Large residual

Regression

Regression Options

Weights:

Confidence level for all intervals:

Type of confidence interval:

Sum of squares for tests:

Box-Cox transformation

log transformation

Optimal lambda

lambda = 2 (natural log)

lambda = 0.5 (square root)

lambda:

Select

Help

OK

Cancel

Current Worksheet: Worksheet 2

Help and Settings

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Regression Analysis: Yi versus Xi

Analysis of Variance

Source	DF	Seq SS	Contribution	Adj SS	Adj MS	F-Value	P-Value
Regression	1	3.588	7.26%	3.588	3.5878	9.24	0.003
Xi	1	3.588	7.26%	3.588	3.5878	9.24	0.003
Error	118	45.818	92.74%	45.818	0.3883		
Lack-of-Fit	19	6.486	13.13%	6.486	0.3414	0.86	0.632
Pure Error	99	39.332	79.61%	39.332	0.3973		
Total	119	49.405	100.00%				

Model Summary

S	R-sq	R-sq(adj)	PRESS	R-sq(pred)
0.623125	7.26%	6.48%	47.6103	3.63%

Coefficients

Term	Coef	SE Coef	99% CI	T-Value	P-Value	VIF
Constant	2.114	0.321	(1.274, 2.954)	6.59	0.000	
Xi	0.0388	0.0128	(0.0054, 0.0723)	3.04	0.003	1.00

Regression Equation

$$Y_i = 2.114 + 0.0388 X_i$$

Fits and Diagnostics for Unusual Observations

Obs	Yi	Fit	SE Fit	99% CI	Resid	Std Resid	Del Resid	HI	Cook's D
2	3.885	2.658	0.148	(2.269, 3.046)	1.227	2.03	2.06	0.0566650	0.12
9	0.500	3.240	0.079	(3.034, 3.446)	-2.740	-4.43	-4.83	0.0160124	0.16
101	1.841	3.085	0.057	(2.936, 3.234)	-1.244	-2.00	-2.03	0.0058651	0.02
102	1.858	2.813	0.103	(2.548, 3.078)	-1.230	-2.00	-2.03	0.0273863	0.06
106	3.716	3.473	0.143	(3.099, 3.847)	0.243	0.40	0.40	0.0526942	0.00
115	1.486	3.318	0.098	(3.060, 3.575)	-1.832	-2.98	-3.08	0.0248782	0.11

Obs DFITS

Obs	DFITS
2	0.503787 R X
9	-0.616775 R
101	-0.166512 R
102	-0.339912 R
106	0.094161 X
115	-0.452302 R

R Large residual

Regression Analysis: Yi versus Xi

Analysis of Variance

Source	DF	Seq SS	Contribution	Adj SS	Adj MS	F-Value	P-Value
Regression	1	3.588	7.26%	3.588	3.5878	9.24	0.003
Xi	1	3.588	7.26%	3.588	3.5878	9.24	0.003
Error	n-2=118	45.818	92.74%	SSE=45.818	MSE=0.3883		
Lack-of-Fit	19	6.486	13.13%	6.486	0.3414	0.86	0.632

Pure Error	99	39.332	79.61%	39.332	0.3973
Total	119	49.405	100.00%		

Model Summary

S	R-sq	R-sq(adj)	PRESS	R-sq(pred)
0.623125	7.26%	6.48%	47.6103	3.63%

Coefficients

Term	Coef	SE Coef	99% CI	T-Value	P-Value	VIF
Constant	2.114	0.321	(1.274, 2.954)	6.59	0.000	
Xi	0.0388	0.0128	(0.0054, 0.0723)	3.04	0.003	1.00

Regression Equation

$$Y_i = 2.114 + 0.0388 X_i$$

$$99\% \text{ C.I for } \beta_1: b_1 - t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) \leq \beta_1 \leq b_1 + t_{(1-\alpha/2, n-2)}s(b_1)$$

$$0.0054 \leq \beta_1 \leq 0.0723$$

Interpret your confidence interval. Does it include zero? No

Why might the director of admissions be interested in whether the confidence interval includes zero?

If the C.I of β_1 include zero, then β_1 can tack zero and $\beta_1 = 0$

b. Test, using the test statistic t^* , whether or not a linear association exists between student's ACT score (X) and GPA at the end of the freshman year (Y). Use a level of significance of 0.01 State the alternatives, decision rule, and conclusion.

$$\alpha = 0.01$$

1. Hypothesis

$$H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

2. Test statistic

$$T_0 = \frac{b_1 - \beta_{10}}{s(b_1)} = \frac{b_1}{s(b_1)} = \frac{0.0388}{0.0128} = 3.04$$

3. Decision: Reject H_0 if $|T_0| > t_{(1-\frac{\alpha}{2}, n-2)}$, $3.04 > t_{(0.995, 118)} = 1.70943$

Then reject H_0

c. What is the P-value of your test in part (b)? How does it support the conclusion reached in part (b)?

p-value= 0.003<0.01, then we reject H_0 .

Q2.5. Refer to Copier maintenance Problem 1.20.

$$n = 45, \sum_{i=1}^{n=45} X_i = 230, \sum_{i=1}^{45} Y_i = 3432, \sum_{i=1}^{45} X_i^2 = 1516, \sum_{i=1}^{45} X_i Y_i = 22660$$

$$SSE = 3416.377$$

a. Estimate the change in the mean service time when the number of copiers serviced increases by one. Use a 90 percent confidence interval. Interpret your confidence interval.

$$90\% \text{ C.I for } \beta_1: b_1 - t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) \leq \beta_1 \leq b_1 + t_{(1-\frac{\alpha}{2}, n-2)}s(b_1)$$

$$\alpha = 1 - 0.9 = 0.1$$

$$b_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i Y_i - \bar{X} Y_i - X_i \bar{Y} + \bar{X} \bar{Y})}{\sum_{i=1}^n (X_i^2 - 2\bar{X} X_i + \bar{X}^2)} = \frac{\sum_{i=1}^n X_i Y_i - n\bar{X}\bar{Y}}{\sum_{i=1}^n X_i^2 - n\bar{X}^2} = \frac{22660 - 45 * 5.1111 * 76.2667}{1516 - 45 * 5.1111^2} = 15.035$$

$$s^2(b_1) = \frac{MSE}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{3416.377 / (45 - 2)}{1516 - 45 * 5.1111^2} = 0.23337$$

$$s(b_1) = \sqrt{0.23337} = 0.48308$$

$$t_{(1-\frac{\alpha}{2}, n-2)} = t_{(0.95, 43)} = 1.68107$$

$$b_1 - t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) = 15.035 - 1.68107 * 0.48308 = 14.222$$

$$b_1 + t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) = 15.035 + 1.68107 * 0.48308 = 15.84709$$

$$14.222 \leq \beta_1 \leq 15.847$$

b. Conduct a t test to determine whether or not there is a linear association between X and Y here; control the α a risk at 0.01. State the alternatives, decision rule, and conclusion. What is the P-value of your test?

$$\alpha = 0.01$$

1. Hypothesis

$$H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

2. Test statistic

$$T_0 = \frac{b_1 - \beta_{10}}{s(b_1)} = \frac{b_1}{s(b_1)} = \frac{15.035}{0.48308} = 31.123$$

3. Decision: Reject H_0 if $|T_0| > t_{(1-\frac{\alpha}{2}, n-2)}$, $31.123 > t_{(0.995, 43)} = 2.695$

Then reject H_0

$$\text{p-value} = 2P(t_{(n-2)} > |T_0|) = 2(1 - P(t_{(n-2)} < 31.123)) = 2(1 - 1)$$

$0.00 < 0.01$, then we reject H_0 .

c. Are your results in parts (a) and (b) consistent? Explain.

Yes, the C.I of β_1 does not include zero, and we reject H_0 .

d. The manufacturer has suggested that the mean required time should not increase by more than 14 minutes for each additional copier that is serviced on a service call. Conduct a test to decide whether this standard is being satisfied by Tri-City. Control the risk of a Type I error at 0.05. State the alternatives, decision rule, and conclusion. What is the P-value of the test?

$$\alpha = 0.05$$

1. Hypothesis

$$H_0: \beta_1 \leq 14$$

$$H_1: \beta_1 > 14$$

2. Test statistic

$$T_0 = \frac{b_1 - \beta_{10}}{s(b_1)} = \frac{b_1 - 14}{s(b_1)} = \frac{15.035 - 14}{0.48308} = 2.143$$

3. Decision: Reject H_0 if $T_0 > t_{(1-\alpha, n-2)}$, $2.143 > t_{(0.95, 43)} = 1.861$

Then reject H_0

$$\text{p-value} = P(t_{(n-2)} > T_0) = (1 - P(t_{(n-2)} < 2.143)) = (1 - 0.981) = 0.019 < 0.05$$

, then we reject H_0 .

Q2.6. Refer to Airfreight breakage Problem 1.21.

$$\bar{X} = 1, \bar{Y} = 14.2, \sum_{i=1}^{n=10} (X_i - \bar{X})(Y_i - \bar{Y}) = 40$$

$$\sum_{i=1}^{10} (X_i - \bar{X})^2 = 10, MSE = 2.2$$

a. Estimate β_1 with a 95 percent confidence interval. Interpret your interval estimate.

$$95\% \text{ C.I for } \beta_1: b_1 - t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) \leq \beta_1 \leq b_1 + t_{(1-\alpha/2, n-2)}s(b_1)$$

$$\alpha = 1 - 0.95 = 0.05$$

$$b_1 = \hat{\beta}_1 = \frac{\sum_{i=1}^{n=120} (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^{n=120} (X_i - \bar{X})^2} = 4$$

$$s^2(b_1) = \frac{MSE}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{2.2}{10} = 0.22$$

$$s(b_1) = \sqrt{0.22} = 0.469$$

$$t_{(1-\frac{\alpha}{2}, n-2)} = t_{(0.975, 8)} = 2.306$$

$$b_1 - t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) = 4 - 2.306 * 0.469 = 2.918$$

$$b_1 + t_{(1-\frac{\alpha}{2}, n-2)}s(b_1) = 4 + 2.306 * 0.469 = 5.081$$

$$2.918 \leq \beta_1 \leq 5.081$$

b. Conduct a t test to decide whether or not there is a linear association between number of times a carton is transferred (X) and number of broken ampules (Y). Use a level of significance of 0.05. State the alternatives, decision rule, and conclusion. What is the P-value of the test?

$$\alpha = 0.05$$

1. Hypothesis

$$H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

2. Test statistic

$$T_0 = \frac{b_1 - \beta_{10}}{s(b_1)} = \frac{b_1}{s(b_1)} = \frac{4}{0.469} = 8.528$$

3. Decision: Reject H_0 if $|T_0| > t_{(1-\frac{\alpha}{2}, n-2)}$, $8.528 > t_{(0.975, 8)} = 2.308$

Then reject H_0

$$p\text{-value} = 2P(t_{(n-2)} > |T_0|) = 2(1 - P(t_{(8)} < 8.528)) = 2(1 - 0.9999)$$

$0.0002 < 0.05$, then we reject H_0 .

Analysis of Variance

Source	DF	Seq SS	Contribution	Adj SS	Adj MS	F-Value	P-Value
Regression	1	160.000	90.09%	160.000	160.000	72.73	0.000
Xi	1	160.000	90.09%	160.000	160.000	72.73	0.000
Error	8	17.600	9.91%	17.600	2.200		
Lack-of-Fit	2	0.933	0.53%	0.933	0.467	0.17	0.849
Pure Error	6	16.667	9.38%	16.667	2.778		
Total	9	177.600	100.00%				

Model Summary

S	R-sq	R-sq(adj)	PRESS	R-sq(pred)
1.48324	90.09%	88.85%	25.8529	85.44%

Coefficients

Term	Coef	SE Coef	95% CI	T-Value	P-Value	VIF
Constant	10.200	0.663	(8.670, 11.730)	15.38	0.000	
X_i	4.000	0.469	(2.918, 5.082)	8.53	0.000	1.00

Regression Equation

$$Y_i = 10.200 + 4.000 X_i$$

H.W:

Q2.7 Refer to Plastic hardness Problem 1.22.

- a. Estimate the change in the mean hardness when the elapsed time increases by one hour. Use a 99 percent confidence interval. Interpret your interval estimate.**
- b. The plastic manufacturer has stated that the mean hardness should increase by 2 Brinell units per hour. Conduct a two-sided test to decide whether this standard is being satisfied; use $\alpha = 0.01$. State the alternatives, decision rule, and conclusion. What is the P-value of the test?**

Chapter 2

We assume that the normal error regression model is applicable. This model is:

$$Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$$

where:

β_0 and β_1 , are parameters

X_i are known constants

ε_i are independent $N(0, \sigma^2)$

$$E(Y_i) = \beta_0 + \beta_1 X_i$$

Sampling Distribution of $\widehat{\beta}_1$

$$\widehat{\beta}_1 = b_1 = \sum_{i=1}^n \frac{(X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$E(\widehat{\beta}_1) = \beta_1$$

$$\sigma^2(\widehat{\beta}_1) = \frac{\sigma^2}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$s^2(\widehat{\beta}_1) = \frac{MSE}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$\frac{b_1 - \beta_1}{s(b_1)} \sim t_{(n-2)}$$

Confidence Interval for β_1

$$P \left[b_1 - t_{(1-\frac{\alpha}{2}, n-2)} s(b_1) \leq \beta_1 \leq b_1 + t_{(1-\alpha/2, n-2)} s(b_1) \right] = 1 - \alpha$$

C.I (1 - α)% for β_1

$$b_1 - t_{(1-\frac{\alpha}{2}, n-2)} s(b_1) \leq \beta_1 \leq b_1 + t_{(1-\alpha/2, n-2)} s(b_1)$$

Tests Concerning β_1

1. Hypothesis		
$H_0: \beta_1 = \beta_{10}$ $H_1: \beta_1 \neq \beta_{10}$	$H_0: \beta_1 = \beta_{10}$ $H_1: \beta_1 > \beta_{10}$	$H_0: \beta_1 = \beta_{10}$ $H_1: \beta_1 < \beta_{10}$
2. Test statistic		
$T_0 = \frac{b_1 - \beta_{10}}{s(b_1)}$		
3. Decision: Reject H_0 if		
$ T_0 > t_{(1-\frac{\alpha}{2}, n-2)}$	$T_0 > t_{(1-\alpha, n-2)}$	$T_0 < t_{(\alpha, n-2)}$
P-value: Reject H_0 if $p - value < \alpha$		
p-value = $2P(t_{(n-2)} > T_0)$	p-value = $P(t_{(n-2)} > T_0)$	p-value = $P(t_{(n-2)} < T_0)$

Sampling Distribution of $\widehat{\beta}_0$

$$\widehat{\beta}_0 = b_0 = \bar{Y} - b_1 \bar{X}$$

$$E(\widehat{\beta}_0) = \beta_0$$

$$\sigma^2(\widehat{\beta}_0) = \sigma^2 \left(\frac{1}{n} + \frac{\bar{X}^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

$$s^2(\widehat{\beta}_0) = MSE \left(\frac{1}{n} + \frac{\bar{X}^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

$$\frac{b_0 - \beta_0}{s(b_0)} \sim t_{(n-2)}$$

Confidence Interval for β_1

$$P \left[b_0 - t_{(1-\frac{\alpha}{2}, n-2)} s(b_0) \leq \beta_0 \leq b_0 + t_{(1-\alpha/2, n-2)} s(b_0) \right] = 1 - \alpha$$

C.I $(1 - \alpha)\%$ for β_0

$$b_0 - t_{(1-\frac{\alpha}{2}, n-2)} s(b_0) \leq \beta_0 \leq b_0 + t_{(1-\alpha/2, n-2)} s(b_0)$$

Tests Concerning β_1

1. Hypothesis		
$H_0: \beta_0 = \beta_{00}$ $H_1: \beta_0 \neq \beta_{00}$	$H_0: \beta_0 = \beta_{00}$ $H_1: \beta_0 > \beta_{00}$	$H_0: \beta_1 = \beta_{00}$ $H_1: \beta_1 < \beta_{00}$
2. Test statistic		
$T_0 = \frac{b_0 - \beta_{00}}{s(b_0)}$		
3. Decision: Reject H_0 if		
$ T_0 > t_{(1-\frac{\alpha}{2}, n-2)}$	$T_0 > t_{(1-\alpha, n-2)}$	$T_0 < t_{(\alpha, n-2)}$
P-value: Reject H_0 if $p - value < \alpha$		
p-value = $2P(t_{(n-2)} > T_0)$	p-value = $P(t_{(n-2)} > T_0)$	p-value = $P(t_{(n-2)} < T_0)$

$$Y_h = b_0 + b_1 X_h$$

ANOVA TABLE

Source of Variation	d.f	SS	MS	F	p-value
Regression	1	$SSR = \sum(\hat{Y}_i - \bar{Y})^2$	$MSR = \frac{SSR}{1}$	$\frac{MSR}{MSE}$	
Error	n-2	$SSE = \sum(Y_i - \hat{Y}_i)^2$	$MSE = \frac{SSE}{n-2}$		
Total	n-1	$SSTo = \sum(Y_i - \bar{Y})^2$			

1. Hypothesis

$$H_0: \beta_1 = 0 \text{ (Non liner)}$$

$$H_1: \beta_1 \neq 0$$

2. Test statistic

$$F^* = \frac{MSR}{MSE}$$

3. Decision: Reject H_0 if

$$F > F_{(1-\alpha, 1, n-2)}$$

P-value: Reject H_0 if $p - value < \alpha$

$$p - value = P(F_{(1, n-2)} > F^*)$$

Q2.6. Refer to Airfreight breakage Problem 1.21.

$$\bar{X} = 1, \bar{Y} = 14.2,$$

$$\sum_{i=1}^{n=10} (X_i - \bar{X})(Y_i - \bar{Y}) = 40, \sum_{i=1}^{10} (X_i - \bar{X})^2 = 10$$

$$\sum_{i=1}^{n=10} (Y_i - \bar{Y})^2 = 177.6, MSE = 2.2$$

$$b_0 = 10.2, b_1 = 4$$

d) A consultant has suggested, on the basis of previous experience, that the mean number of broken ampules should not exceed 9.0 when no transfers are made. Conduct an appropriate test, using $\alpha = 0.025$. State the alternatives, decision rule, and conclusion. What is the P-value of the test?

$$\alpha = 0.025$$

1. Hypothesis

$$H_0: \beta_0 \leq 9$$

$$H_1: \beta_0 > 9$$

2. Test statistic

$$T_0 = \frac{b_0 - \beta_{00}}{s(b_0)} = \frac{10.2 - 9}{0.6633} = 1.809$$

$$s^2(\widehat{\beta}_0) = MSE \left(\frac{1}{n} + \frac{\bar{X}^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right) = 2.2 \left(\frac{1}{10} + \frac{1^2}{10} \right) = 0.44$$

$$s(b_0) = 0.6633$$

3. Decision: Reject H_0 if $T_0 > t_{(1-\alpha, n-2)}$,

$$1.809 \not> t_{(0.975, 8)} = 2.306$$

Then not reject H_0

$$\text{p-value} = P(t_{(n-2)} > T_0) = (1 - P(t_{(n-2)} < 1.809)) = (1 - 0.945) = 0.055 \not< 0.025$$

, then we not reject H_0 .

at $\alpha = 0.05$

$$b_0 - t_{(1-\frac{\alpha}{2}, n-2)}s(b_0) \leq \beta_0 \leq b_0 + t_{(1-\alpha/2, n-2)}s(b_0)$$

$$t_{(1-\frac{\alpha}{2}, n-2)} = t_{(0.975, 8)} = 2.306$$

$$10.2 - 2.306 * 0.6633 \leq \beta_0 \leq 10.2 + 2.306 * 0.6633$$

$$8.76 \leq \beta_0 \leq 11.728$$

Analysis of Variance

Source	DF	Seq SS	Contribution	Adj	SS Adj	MS	F-Value	P-Value
Regression	1	160.000	90.09%	160.000	160.000	160.000	72.73	0.000
Xi	1	160.000	90.09%	160.000	160.000			
Error	8	17.600	9.91%	17.600		2.200		
Total	9	177.600						100.00%

Coefficients

Term	Coef	SE Coef	95% CI	T-Value	P-Value	VIF
Constant	10.200	0.663	(8.670, 11.730)	15.38	0.000	
Xi	4.000	0.469	(2.918, 5.082)	8.53	0.000	1.00

Regression Equation

$$Y_i = 10.200 + 4.000 X_i$$

Q2.25. Refer to Airfreight breakage Problem 1.21.

a. Set up the ANOVA table. Which elements are additive?

b. Conduct an F test to decide whether or not there is a linear association between the number of times a carton is transferred and the number of broken ampules; control the α risk at 0.05. State the alternatives, decision rule, and conclusion.

c. Obtain the t^* statistic for the test in part (b) and demonstrate numerically its equivalence to the F^* statistic obtained in part (b).

$\bar{X} = 1, \bar{Y} = 14.2, \sum_{i=1}^{n=10} (X_i - \bar{X})(Y_i - \bar{Y}) = 40, \sum_{i=1}^{10} (X_i - \bar{X})^2 = 10$

$\sum_{i=1}^{n=10} (Y_i - \bar{Y})^2 = 177.6, \quad MSE = 2.2, \quad b_0 = 10.2, \quad b_1 = 4$

X_i	Y_i	$(X_i - \bar{X})$	$(Y_i - \bar{Y})$	$(X_i - \bar{X})^2$	$(X_i - \bar{X}) * (Y_i - \bar{Y})$	$(Y_i - \bar{Y})^2$	\hat{Y}_i	$(Y_i - \hat{Y}_i)^2$
1	16	0	1.8	0	0	3.24	14.2	3.24
0	9	-1	-5.2	5.2	1	27.04	10.2	1.44
2	17	1	2.8	2.8	1	7.84	18.2	1.44
0	12	-1	-2.2	2.2	1	4.84	10.2	3.24
3	22	2	7.8	15.6	4	60.84	22.2	0.04
1	13	0	-1.2	0	0	1.44	14.2	1.44
0	8	-1	-6.2	6.2	1	38.44	10.2	4.84
1	15	0	0.8	0	0	0.64	14.2	0.64
2	19	1	4.8	4.8	1	23.04	18.2	0.64
0	11	-1	-3.2	3.2	1	10.24	10.2	0.64
10	142	0	0	40	10	177.6	142	17.6

$$\sum (Y_i - \hat{Y}_i)^2 = 17.6$$

ANOVA TABLE

Source of Variation	d.f	SS	MS	F	p-value
Regression	1	SSR=177.6 - 17.6 = 160	MSR = 160	$\frac{160}{2.2} = 72.72$	0.00
Error	8	SSE=17.6	$MSE = \frac{17.6}{8} = 2.2$		
Total	9	SSTo= 177.6			

$$\alpha = 0.05$$

1. Hypothesis

$$H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

2. Test statistic

$$F^* = 72.72$$

3. Decision: Reject H_0 if $F^* > F_{(1-\alpha, 1, n-2)}$, $72.72 > F_{(0.95, 1, 8)} = 5.31$

Then reject H_0

$$p\text{-value} = P(F_{(1, n-2)} > F^*) = (1 - P(F_{(1, 8)} < 72.72)) = (1 - 0.9999) = 0.0001 < 0.05$$

, then we reject H_0 .

Analysis of Variance

Source	DF	Adj SS	Adj MS	F-Value	P-Value
Regression	1	160.000	160.000	72.73	0.000
Xi	1	160.000	160.000	72.73	0.000
Error	8	17.600	2.200		
Lack-of-Fit	2	0.933	0.467	0.17	0.849
Pure Error	6	16.667	2.778		
Total	9	177.600			

$$t^* = 8.528, (t^*)^2 = (8.528)^2 = 72.72 = F^*$$

Q2.26. Refer to Plastic hardness Problem 1.22.

a. Set up the ANOVA table.

b. Test by means of an F test whether or not there is a linear association between the hardness of the plastic and the elapsed time. Use $\alpha = .01$. State the alternatives, decision rule, and conclusion.

Analysis of Variance

Source	DF	Adj SS	Adj MS	F-Value	P-Value
Regression	1	5297.51	5297.51	506.51	0.000
Xi	1	5297.51	5297.51	506.51	0.000
Error	14	146.43	10.46		
Lack-of-Fit	2	17.67	8.84	0.82	0.462
Pure Error	12	128.75	10.73		
Total	15	5443.94			

$$\alpha = 0.01$$

1. Hypothesis

$$H_0: \beta_1 = 0$$

$$H_1: \beta_1 \neq 0$$

2. Test statistic

$$F^* = 506.51$$

3. Decision:

$$p\text{-value} = 0.000 < 0.01$$

, then we reject H_0 .

Prove that

$$Q = \sum_{i=1}^n \epsilon_i^2 = \sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)^2$$

$$\frac{\partial Q}{\partial \beta_0} = 0, \frac{\partial Q}{\partial \beta_1} = 0$$

$$\frac{\partial Q}{\partial \beta_0} = 2 \sum_{i=1}^n (Y_i - b_0 - b_1 X_i) (-1) = 0$$

$$\sum_{i=1}^n (Y_i - b_0 - b_1 X_i) = 0$$

$$\sum_{i=1}^n (Y_i) - \sum_{i=1}^n (b_0) - \sum_{i=1}^n (b_1 X_i) = 0$$

$$\sum_{i=1}^n (Y_i) - nb_0 - b_1 \sum_{i=1}^n (X_i) = 0$$

$$\sum_{i=1}^n (Y_i) = nb_0 + b_1 \sum_{i=1}^n (X_i) \rightarrow (1)$$

$$\frac{\partial Q}{\partial \beta_0} = 2 \sum_{i=1}^n [(Y_i - b_0 - b_1 X_i)(-X_i)] = 0$$

$$\sum_{i=1}^n (Y_i X_i - b_0 X_i - b_1 X_i^2) = 0$$

$$\sum_{i=1}^n (Y_i X_i) - b_0 \sum_{i=1}^n (X_i) - b_1 \sum_{i=1}^n (X_i^2) = 0$$

$$\sum_{i=1}^n (Y_i X_i) = b_0 \sum_{i=1}^n (X_i) + b_1 \sum_{i=1}^n (X_i^2) \rightarrow (2)$$

By solving 1 and 2 together

$$\sum_{i=1}^n (Y_i) = n b_0 + b_1 \sum_{i=1}^n (X_i)$$

$$\sum_{i=1}^n (Y_i X_i) = b_0 \sum_{i=1}^n (X_i) + b_1 \sum_{i=1}^n (X_i^2)$$

From 1

$$\bar{Y} = b_0 + b_1 \bar{X}$$

$$b_0 = \bar{Y} - b_1 \bar{X}$$

$$\sum_{i=1}^n (Y_i X_i) = (\bar{Y} - b_1 \bar{X}) \sum_{i=1}^n (X_i) + b_1 \sum_{i=1}^n (X_i^2)$$

$$\sum_{i=1}^n (Y_i X_i) = \bar{Y} \sum_{i=1}^n (X_i) + b_1 \left[\sum_{i=1}^n (X_i^2) - \bar{X} \sum_{i=1}^n (X_i) \right]$$

$$\sum_{i=1}^n (Y_i X_i) - \bar{Y} \sum_{i=1}^n (X_i) = b_1 \left[\sum_{i=1}^n (X_i^2) - \bar{X} \sum_{i=1}^n (X_i) \right]$$

$$\sum_{i=1}^n (Y_i X_i) - n\bar{Y}\bar{X} = b_1 \left[\sum_{i=1}^n (X_i^2) - n\bar{X}^2 \right]$$

$$b_1 = \frac{\sum_{i=1}^n (Y_i X_i) - n\bar{Y}\bar{X}}{\left[\sum_{i=1}^n (X_i^2) - n\bar{X}^2 \right]} = \sum_{i=1}^n \frac{(X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$\left(\sum_{i=1}^n X_i \right)^2 \neq \sum_{i=1}^n (X_i^2)$$

$$\sum_{i=1}^n \frac{(X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$\left(\sum_{i=1}^n (X_i - \bar{X})^2 \right)^2 \neq \sum_{i=1}^n (X_i - \bar{X})^4$$

$$\sum_{i=1}^n e_i = 0$$

$$\sum_{i=1}^n e_i = \sum_{i=1}^n (Y_i - \hat{Y}_i) = \sum_{i=1}^n (Y_i - b_0 - b_1 X_i)$$

$$= \sum_{i=1}^n (Y_i - \bar{Y} + b_1 \bar{X} - b_1 X_i) = \sum_{i=1}^n (Y_i - \bar{Y}) - b_1 \sum_{i=1}^n (X_i - \bar{X}) = 0 - 0 = 0$$

$$b_0 = \bar{Y} - b_1\bar{X}$$

$$\begin{aligned}
\sum_{i=1}^n e_i X_i = 0, \quad \sum_{i=1}^n e_i X_i &= \sum_{i=1}^n [(Y_i - b_0 - b_1 X_i) X_i] = \sum_{i=1}^n (Y_i X_i - b_0 X_i - b_1 X_i^2) = \sum_{i=1}^n (Y_i X_i) - b_0 \sum_{i=1}^n (X_i) - b_1 \sum_{i=1}^n (X_i^2) \\
&= \sum_{i=1}^n (Y_i X_i) - (\bar{Y} - b_1 \bar{X}) \sum_{i=1}^n (X_i) - b_1 \sum_{i=1}^n (X_i^2) = \sum_{i=1}^n (Y_i X_i) - \bar{Y} \sum_{i=1}^n (X_i) + b_1 \bar{X} \sum_{i=1}^n (X_i) - b_1 \sum_{i=1}^n (X_i^2) \\
&= \sum_{i=1}^n (Y_i X_i) - n\bar{Y}\bar{X} + nb_1\bar{X}^2 - b_1 \sum_{i=1}^n (X_i^2) = \left[\sum_{i=1}^n (Y_i X_i) - n\bar{Y}\bar{X} \right] - b_1 \left[\sum_{i=1}^n (X_i^2) - n\bar{X}^2 \right] \\
&= \sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y}) - b_1 \sum_{i=1}^n (X_i - \bar{X})^2 = \sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y}) - \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2} \sum_{i=1}^n (X_i - \bar{X})^2 = 0 \\
\sum_{i=1}^n \hat{Y}_i &= \sum_{i=1}^n (b_0 + b_1 X_i) = \sum_{i=1}^n (\bar{Y} - b_1 \bar{X} + b_1 X_i) = n\bar{Y} + b_1 \sum_{i=1}^n (X_i - \bar{X}) = \sum_{i=1}^n Y_i
\end{aligned}$$

$$Var\left(\frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}\right)$$

$$\frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})Y_i - \bar{Y} \sum_{i=1}^n (X_i - \bar{X})}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})Y_i}{\sum_{i=1}^n (X_i - \bar{X})^2}$$

$$\begin{aligned} \text{Var}\left(\frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2}\right) &= \text{Var}\left(\frac{\sum_{i=1}^n (X_i - \bar{X})Y_i}{\sum_{i=1}^n (X_i - \bar{X})^2}\right) = \text{Var}\left(\sum_{i=1}^n \frac{(X_i - \bar{X})}{\sum_{i=1}^n (X_i - \bar{X})^2} Y_i\right) = \sum_{i=1}^n \left(\frac{(X_i - \bar{X})}{\sum_{i=1}^n (X_i - \bar{X})^2}\right)^2 \text{Var}(Y_i) \\ &= \sum_{i=1}^n \frac{(X_i - \bar{X})^2}{(\sum_{i=1}^n (X_i - \bar{X})^2)^2} \sigma^2 = \sigma^2 \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{(\sum_{i=1}^n (X_i - \bar{X})^2)^2} = \frac{\sigma^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \end{aligned}$$

Prove that SSTo=SSR+SSE.

$$\begin{aligned} \text{L.H.S} = \text{SSTo} &= \sum_{i=1}^n (Y_i - \bar{Y})^2 = \sum_{i=1}^n (Y_i - \hat{Y}_i + \hat{Y}_i - \bar{Y})^2 = \sum_{i=1}^n \left((Y_i - \hat{Y}_i) + (\hat{Y}_i - \bar{Y}) \right)^2 = \sum_{i=1}^n \left[(Y_i - \hat{Y}_i)^2 + (\hat{Y}_i - \bar{Y})^2 + \right. \\ &\left. 2(Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) \right] \end{aligned}$$

$$= \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 + \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 + 2 \sum_{i=1}^n (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y})$$

$$\because (Y_i - \hat{Y}_i) = e_i$$

$$\sum_{i=1}^n (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) = \sum_{i=1}^n e_i(\hat{Y}_i - \bar{Y}) = \sum_{i=1}^n e_i \hat{Y}_i - \bar{Y} \sum_{i=1}^n e_i$$

$$\because \sum_{i=1}^n e_i \hat{Y}_i = \sum_{i=1}^n e_i = 0$$

Then

$$\sum_{i=1}^n (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) = 0$$

Then

$$SST_o = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2 + \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2$$

$$\sum_{i=1}^n (Y_i - \hat{Y}_i)^2 = SSE \quad \& \quad \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 = SSR$$

$$SST_o = SSR + SSE = L.H.S$$

Chapter 2

2.13 Refer to Grade point average.

Calculate R^2 . What proportion of the variation in Y is accounted for by introducing X into the regression model? From page 98

$$\bar{X} = 24.725, \sum_{i=1}^{n=120} (X_i - \bar{X})^2 = 2379.925$$

Analysis of Variance

Source	DF	Adj SS	Adj MS	F-Value	P-Value
Regression	1	SSR=3.588	3.5878	9.24	0.003
Xi	1	3.588	3.5878	9.24	0.003
Error	118	SSE=45.818	MSE=0.3883		
Lack-of-Fit	19	6.486	0.3414	0.86	0.632
Pure Error	99	39.332	0.3973		
Total	119	SSTo=49.405			

Model Summary

S	R-sq	R-sq(adj)	R-sq(pred)
0.623125	7.26%	6.48%	3.63%

$$R^2 = \frac{SSR}{SSTo} = \frac{3.588}{49.405} = 0.0726$$

$$R^2 = 1 - \frac{SSE}{SSTo} = 1 - \frac{45.818}{49.405} = 1 - 0.9274 = 0.0726$$

This means that 7.26% of change in the mean freshman OPA for students is by ACT test score

a. Obtain a 95 percent interval estimate of the mean freshman OPA for students whose ACT test score is 28. Interpret your confidence interval.
From page 76- to 79

$$\hat{Y}_h = b_0 + b_1 X_h$$

$$s^2(\hat{Y}_h) = MSE \left(\frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

$$\hat{Y}_h \pm t \left(1 - \frac{\alpha}{2}; n - 2 \right) s(\hat{Y}_h)$$

$$\alpha = 0.05, \frac{\alpha}{2} = 0.025$$

$$\text{At } X_h = 28$$

$$\hat{Y}_h = 2.114 + 0.0388 (28) = 3.2012$$

$$s^2(\hat{Y}_h) = MSE \left(\frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

$$s^2(\hat{Y}_h) = MSE \left(\frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right) = \mathbf{0.3883} \left(\frac{1}{120} + \frac{(28 - 24.725)^2}{2379.925} \right) = 0.004986$$

$$s(\hat{Y}_h) = \sqrt{0.007776} = 0.0706$$

$$t \left(1 - \frac{\alpha}{2}; n - 2 \right) = t(0.975; 118) = 1.9807$$

$$3.22012 \pm 1.9807(0.0706)$$

$$3.0614 < E(Y_h) < 3.3410$$

b. Mary Jones obtained a score of 28 on the entrance test. **Predict** her freshman OPA-using a 95 percent prediction interval. Interpret your prediction interval.

$$s^2(\widehat{Y}_{new}) = MSE \left(1 + \frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right)$$

$$\widehat{Y}_h \pm t \left(1 - \frac{\alpha}{2}; n - 2 \right) s(\widehat{Y}_{new})$$

$$s^2(\widehat{Y}_{new}) = MSE \left(1 + \frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right) = \mathbf{0.3883} \left(1 + \frac{1}{120} + \frac{(28 - 24.725)^2}{2379.925} \right) = 0.39328$$

$$s(\widehat{Y}_{new}) = 0.6271$$

$$3.22012 \pm 1.9807(0.6271)$$

$$1.9594 < Y_{h(new)} < 4.4430$$

c. Is the prediction interval in part (b) wider than the confidence interval in part (a)? Should it be?

هل فترة الثقة للتنبؤ في الجزء (ب) أوسع من فترة الثقة في الجزء (أ)؟ هل يجب أن تكون؟

Yes, Yes

2.15. Refer to Airfreight breakage Problem 1.21.

$$\bar{X} = 1, \quad \sum_{i=1}^{10} (X_i - \bar{X})^2 = 10$$

ANOVA TABLE

Source of Variation	d.f	SS	MS	F	p-value
Regression	1	SSR=160	$MSR = 160$	72.72	0.00
Error	8	SSE=17.6	$MSE = 2.2$		
Total	9	SSTo= 177.6			

a. Because of changes in airline routes, shipments may have to be transferred more frequently than in the past. Estimate the mean breakage for the following numbers of transfers: $X = 2, 4$. Use separate 99 percent confidence intervals. Interpret your results.

At $X_h = 2$

$$\hat{Y}_h = 10.2 + 4(2) = 18.2$$

$$s^2(\hat{Y}_h) = MSE \left(\frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right) = 2.2 \left(\frac{1}{10} + \frac{(2 - 1)^2}{10} \right) = 0.44$$

$$s(\hat{Y}_h) = \sqrt{0.44} = 0.6633$$

$$t \left(1 - \frac{\alpha}{2}; n - 2 \right) = t(0.995; 8) = 3.355$$

$$18.2 \pm 3.355(0.6633)$$

$$15.976 < E(Y_h) < 20.424$$

At $X_h = 4$

$$\hat{Y}_h = 10.2 + 4(4) = 26.2$$

$$s^2(\hat{Y}_h) = MSE \left(\frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right) = 2.2 \left(\frac{1}{10} + \frac{(4 - 1)^2}{10} \right) = 2.2$$

$$s(\hat{Y}_h) = \sqrt{2.2} = 1.483$$

$$t \left(1 - \frac{\alpha}{2}; n - 2 \right) = t(0.995; 8) = 3.355$$

$$26.2 \pm 3.355(1.483)$$

$$12.748 < E(Y_h) < 23.652$$

We conclude that the mean number of ampules found to be broken upon arrival when 2 transfers from one aircraft to another over the shipment route of 2 are produced is somewhere between 15.976 and 20.424 ampules

أن متوسط عدد أمبولات وجدت منكسره عند وصولهم عندما تم نقله عبر 2 مرات من طائرة واحدة إلى آخر عبر مسار الشحنه , بين 15.976 و 20.424 أمبوله.

We conclude that the mean number of ampules found to be broken upon arrival when 4 transfers from one aircraft to another over the shipment route are produced is somewhere between 12.748 and 23.652 ampules.

أن متوسط عدد أمبولات وجدت منكسره عند وصولهم عندما تم نقله عبر 4 مرات من طائرة واحدة إلى آخر عبر مسار الشحنه , بين 12.748 و 23.652 أمبوله.

b. The next shipment will entail **two** transfers. Obtain a 99 percent **prediction** interval for the number of broken ampules for this shipment. Interpret your prediction interval.

$$s^2(\widehat{Y}_{new}) = MSE \left(1 + \frac{1}{n} + \frac{(X_h - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \right) = 2.2 \left(1 + \frac{1}{10} + \frac{(2 - 1)^2}{10} \right) = 2.64$$

$$s(\hat{Y}_h) = \sqrt{2.64} = 1.6248$$

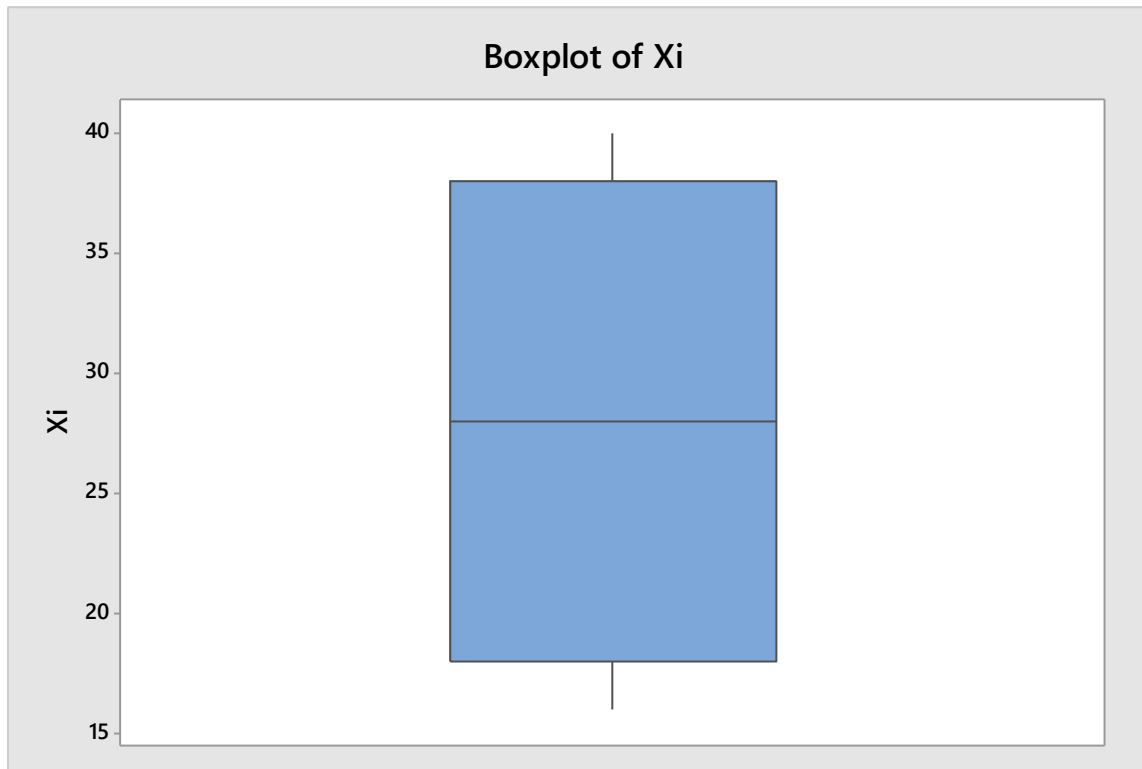
$$18.2 \pm 3.355(1.6248)$$

$$12.748 < Y_{h(new)} < 23.652$$

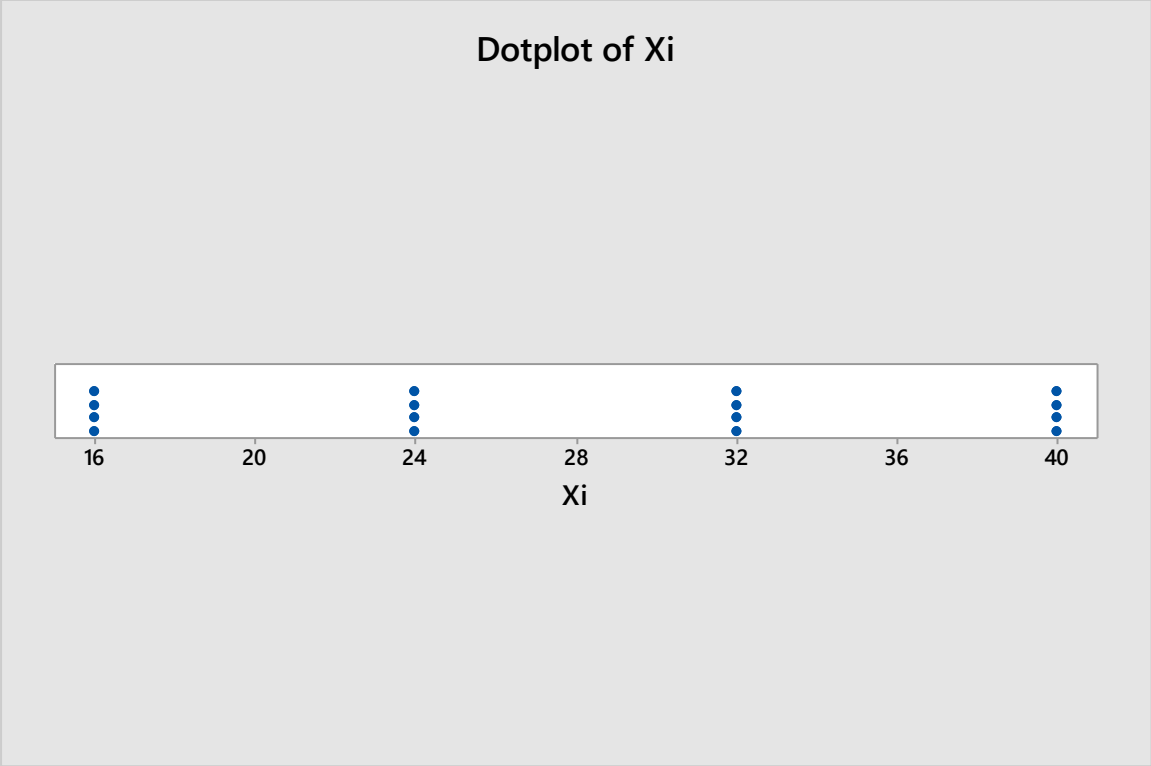
With confidence coefficient 0.99, we predict that the mean number of ampules found to be broken upon arrival when 2 transfers from one aircraft to another over the shipment route of 2 are produced is somewhere between 12.748 and 23.652 ampules.

Refer to Plastic hardness.

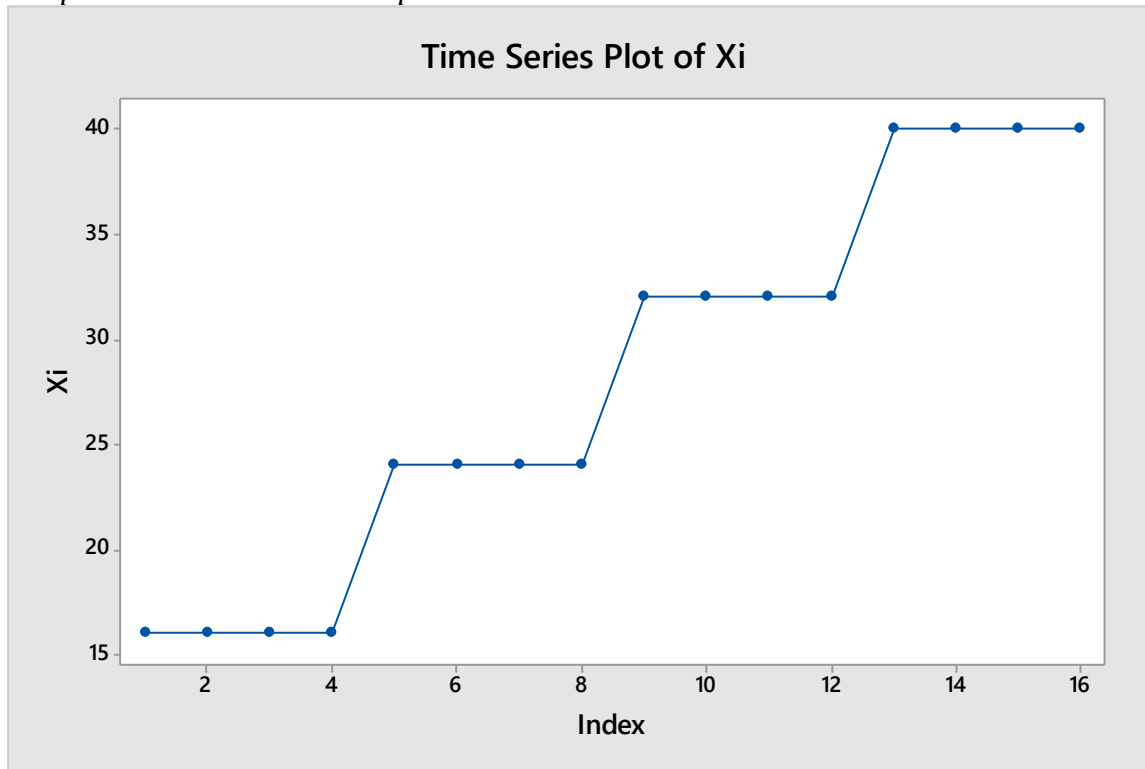
Graph → *Boxplot* → *simple* → *X* → *ok*



Graph → Dotplot → simple → X → ok

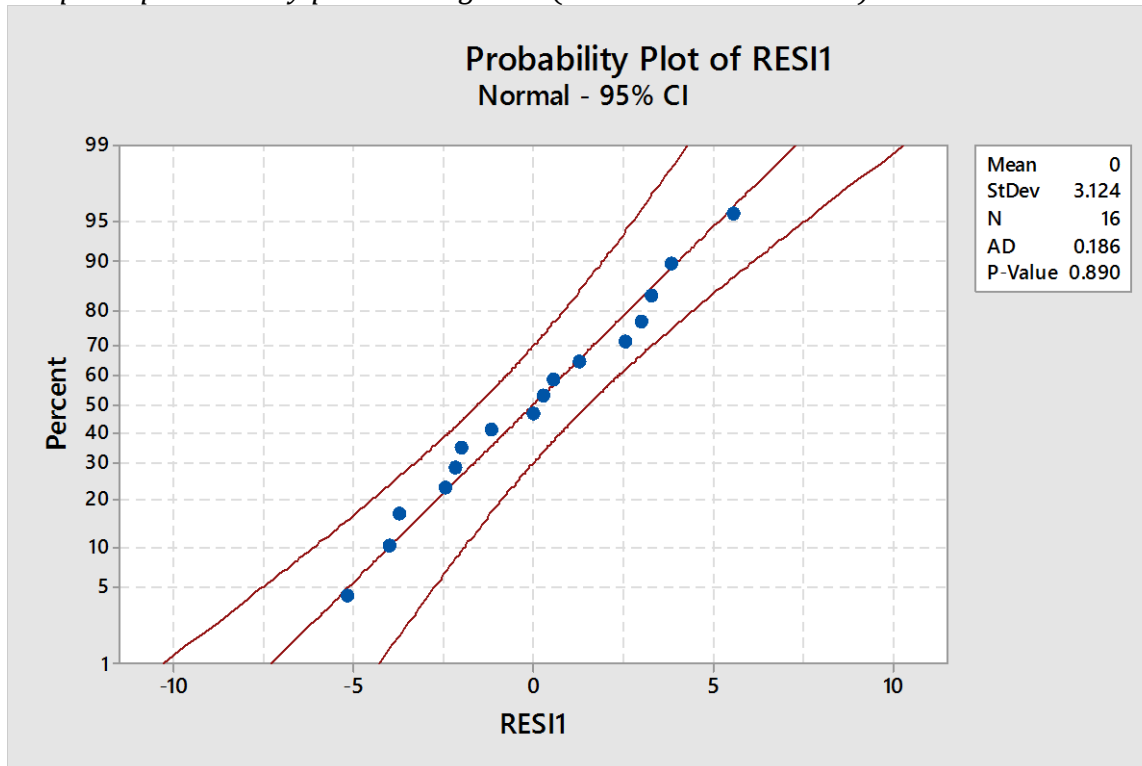


Graph \rightarrow time series \rightarrow simple $\rightarrow X \rightarrow ok$



For test normality of residuals

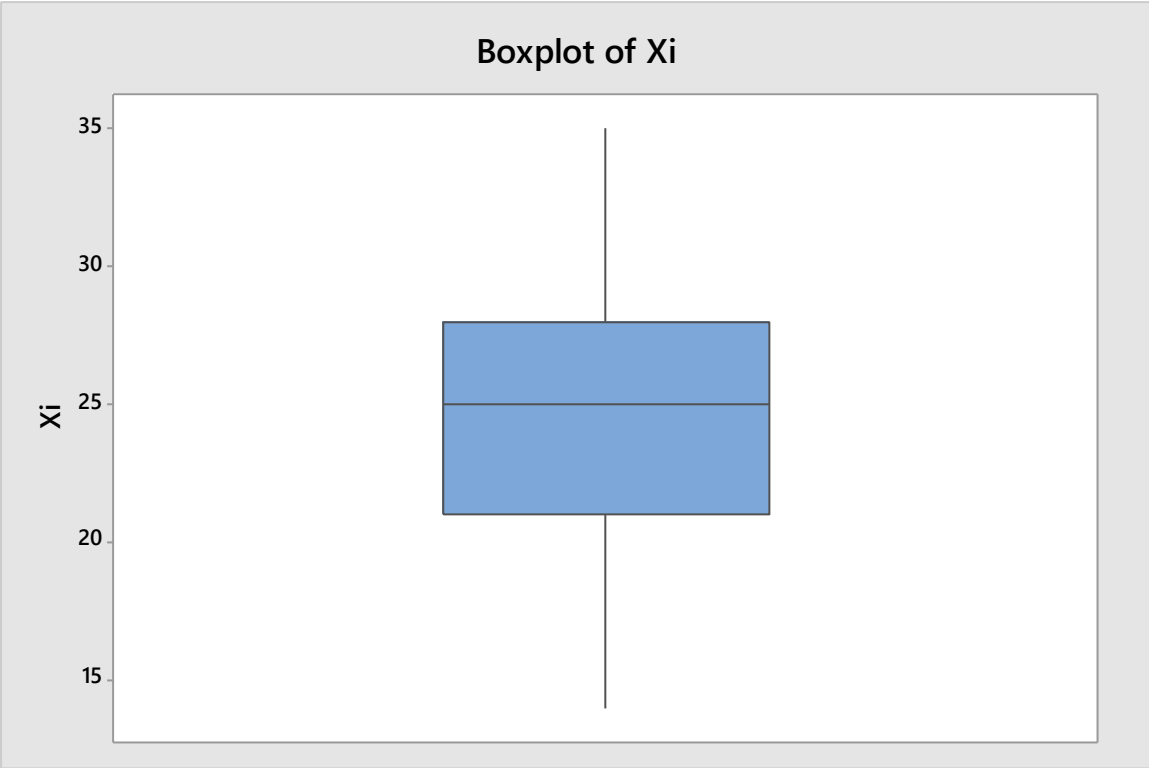
Graph → probability plot → single → (distribution Normal) → X → ok



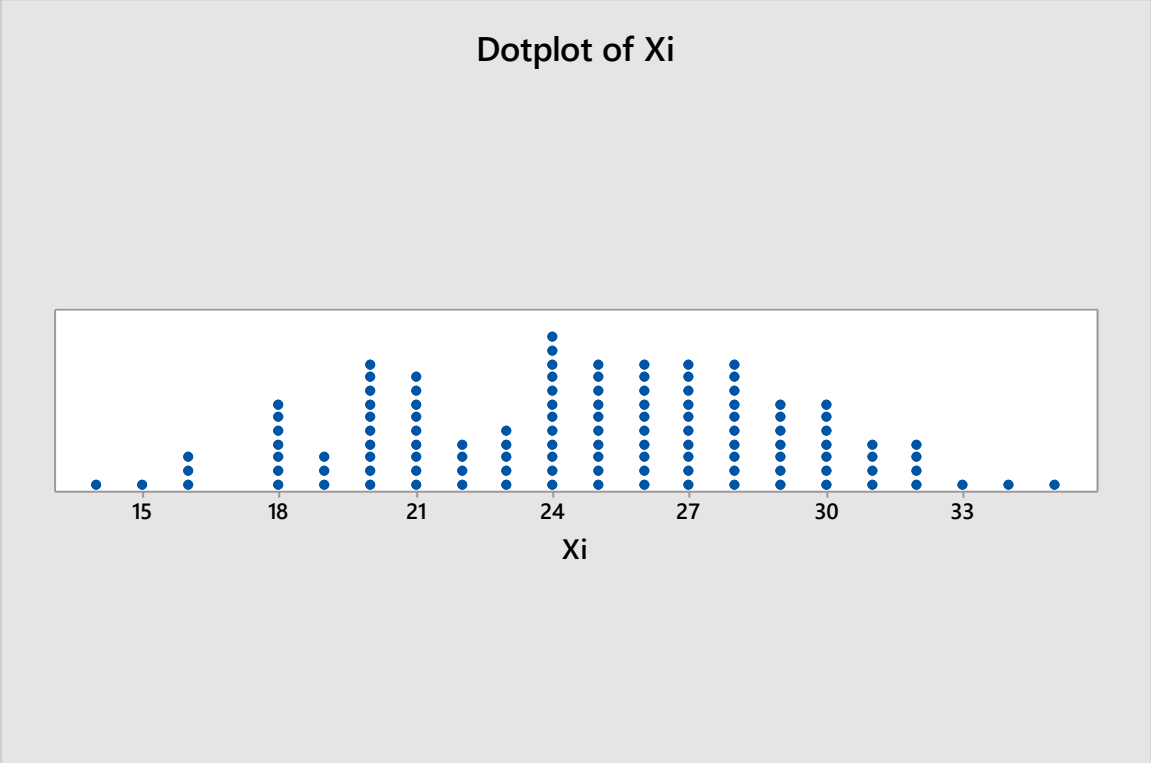
If p-value > 0.05 , then it is normal

Refer to Grade point average.

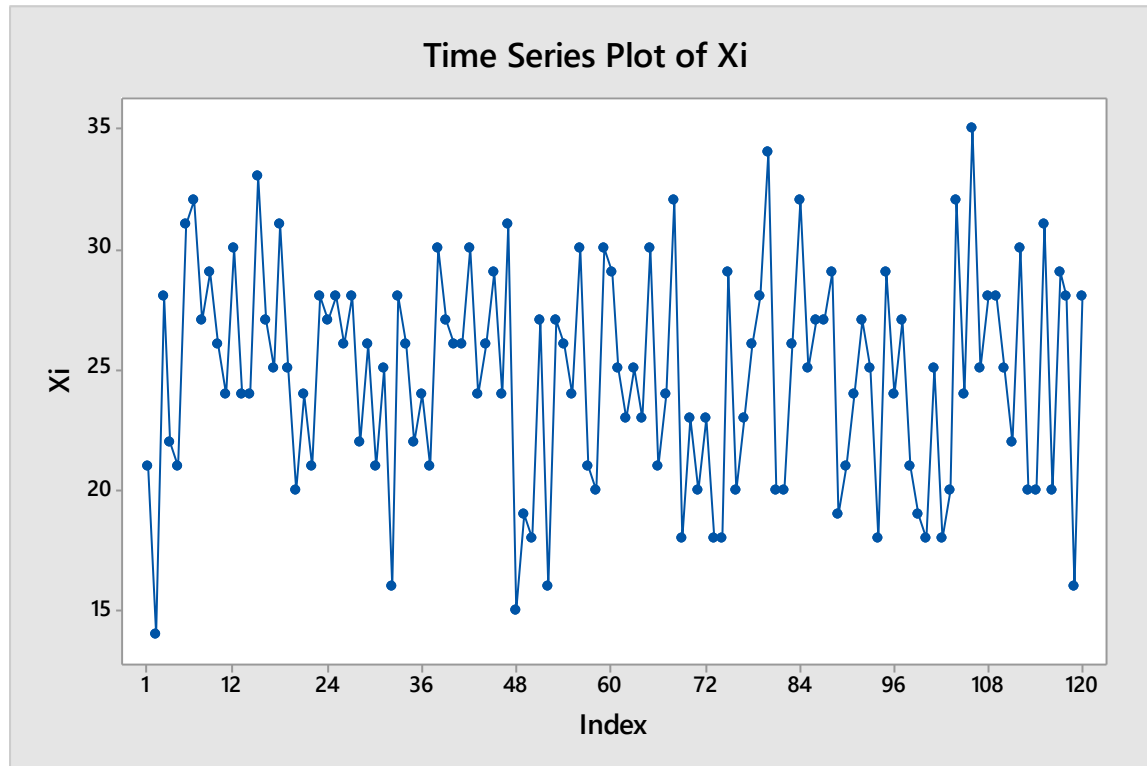
Graph \rightarrow Boxplot \rightarrow simple $\rightarrow X \rightarrow ok$



Graph → Dotplot → simple → X → ok

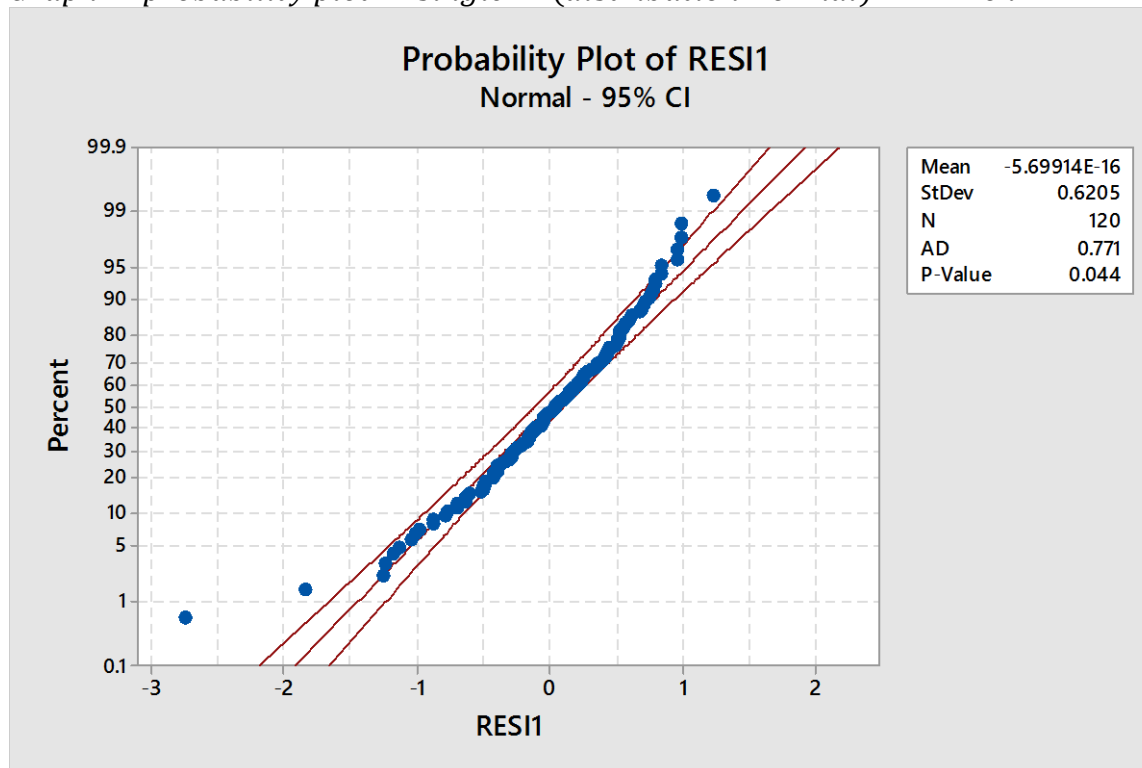


Graph \rightarrow time series \rightarrow simple $\rightarrow X \rightarrow ok$



For test normality of residuals

Graph \rightarrow probability plot \rightarrow single \rightarrow (distribution Normal) \rightarrow X \rightarrow ok



At 0.01 it is normal but at 0.05 , it is not normal

Chapter 5

Q5.1. For the matrices below, obtain (1) $A + B$, (2) $A - B$, (3) AC , (4) AB' , (5) $B'A$.

$$A = \begin{bmatrix} 1 & 4 \\ 2 & 6 \\ 3 & 8 \end{bmatrix} \quad B = \begin{bmatrix} 1 & 3 \\ 1 & 4 \\ 2 & 5 \end{bmatrix} \quad C = \begin{bmatrix} 3 & 8 & 1 \\ 5 & 4 & 0 \end{bmatrix}$$

Solution:

$$A + B = \begin{bmatrix} 1 & 4 \\ 2 & 6 \\ 3 & 8 \end{bmatrix} + \begin{bmatrix} 1 & 3 \\ 1 & 4 \\ 2 & 5 \end{bmatrix} = \begin{bmatrix} 1+1 & 4+3 \\ 2+1 & 6+4 \\ 3+2 & 8+5 \end{bmatrix} = \begin{bmatrix} 2 & 7 \\ 3 & 10 \\ 5 & 13 \end{bmatrix}$$

$$A - B = \begin{bmatrix} 1 & 4 \\ 2 & 6 \\ 3 & 8 \end{bmatrix} - \begin{bmatrix} 1 & 3 \\ 1 & 4 \\ 2 & 5 \end{bmatrix} = \begin{bmatrix} 1-1 & 4-3 \\ 2-1 & 6-4 \\ 3-2 & 8-5 \end{bmatrix} = \begin{bmatrix} 0 & 1 \\ 1 & 2 \\ 1 & 3 \end{bmatrix}$$

$$AC = \begin{bmatrix} 1 & 4 \\ 2 & 6 \\ 3 & 8 \end{bmatrix} \begin{bmatrix} 3 & 8 & 1 \\ 5 & 4 & 0 \end{bmatrix} = \begin{bmatrix} 1*3+4*5 & 1*8+4*4 & 1*1+4*0 \\ 2*3+6*5 & 2*8+6*4 & 2*1+6*0 \\ 3*3+8*5 & 3*8+8*4 & 3*1+8*0 \end{bmatrix} = \begin{bmatrix} 23 & 24 & 1 \\ 36 & 40 & 2 \\ 49 & 56 & 3 \end{bmatrix}$$

$$B' = \begin{bmatrix} 1 & 1 & 2 \\ 3 & 4 & 5 \end{bmatrix}$$

$$(A_{3 \times 2} B'_{2 \times 3})_{3 \times 3} = \begin{bmatrix} 1 & 4 \\ 2 & 6 \\ 3 & 8 \end{bmatrix} \begin{bmatrix} 1 & 1 & 2 \\ 3 & 4 & 5 \end{bmatrix} = \begin{bmatrix} 1*1+4*3 & 1*1+4*4 & 1*2+4*5 \\ 2*1+6*3 & 2*1+6*4 & 2*2+6*5 \\ 3*1+8*3 & 3*1+8*4 & 3*2+8*5 \end{bmatrix} = \begin{bmatrix} 13 & 17 & 22 \\ 20 & 26 & 34 \\ 27 & 35 & 46 \end{bmatrix}$$

$$(B'_{2 \times 3} A_{3 \times 2})_{2 \times 2} = \begin{bmatrix} 1 & 1 & 2 \\ 3 & 4 & 5 \end{bmatrix} \begin{bmatrix} 1 & 4 \\ 2 & 6 \\ 3 & 8 \end{bmatrix} = \begin{bmatrix} 1*1+1*2+2*3 & 1*4+1*6+2*8 \\ 3*1+4*2+5*3 & 3*4+4*6+5*8 \end{bmatrix} = \begin{bmatrix} 9 & 26 \\ 26 & 76 \end{bmatrix}$$

Q5.4. Flavor deterioration. The results shown below were obtained in a small-scale experiment to study the relation between °F of storage temperature (X) and number of weeks before flavour deterioration of a food product begins to occur (Y).

i	1	2	3	4	5
X_i	8	4	0	-4	-8
Y_i	7.8	9.0	10.2	11.0	11.7

Assume that first-order regression model (2.1) is applicable. Using matrix methods, find (1) $Y'Y$, (2) $X'X$, (3) $X'Y$.

$$Y_{n \times 1} = X_{n \times 2} B_{2 \times 1} + \epsilon_{n \times 1}$$

$$E(Y_{n \times 1}) = X_{n \times 2} B_{2 \times 1}$$

$$B_{2 \times 1} = (X'X)^{-1} X'Y$$

$$V(B) = MSE(X'X)^{-1}$$

$$MSE = \frac{e'e}{n-2}$$

	C1	C2	C3
1	8	7.8	1
2	4	9.0	1
3	0	10.2	1
4	-4	11.0	1
5	-8	11.7	1

$$X = \begin{bmatrix} 1 & 8 \\ 1 & 4 \\ 1 & 0 \\ 1 & -4 \\ 1 & -8 \end{bmatrix}, \quad Y = \begin{bmatrix} 7.8 \\ 9.0 \\ 10.2 \\ 11.0 \\ 11.7 \end{bmatrix}, \quad X' = \begin{bmatrix} 1 & 1 & 1 & 1 & 1 \\ 8 & 4 & 0 & -4 & -8 \end{bmatrix}$$

$$(\mathbf{X}'\mathbf{X})_{2 \times 2} = \begin{bmatrix} n & \sum X_i \\ \sum X_i & \sum X_i^2 \end{bmatrix} = \begin{bmatrix} 1 & 1 & 1 & 1 & 1 \\ 8 & 4 & 0 & -4 & -8 \end{bmatrix} \begin{bmatrix} 1 & 8 \\ 1 & 4 \\ 1 & 0 \\ 1 & -4 \\ 1 & -8 \end{bmatrix} = \begin{bmatrix} 1+1+1+1+1 & 8+4+0-4-8 \\ 8+4+0-4-8 & 8*8+4*4+0*0+-4*-4+-8*-8 \end{bmatrix}$$

$$= \begin{bmatrix} 5 & 0 \\ 0 & 160 \end{bmatrix}$$

$$(\mathbf{X}'\mathbf{X})^{-1} = \frac{1}{\Delta} \begin{bmatrix} \sum X_i^2 & -\sum X_i \\ -\sum X_i & n \end{bmatrix}$$

$$\Delta = n \sum X_i^2 - \sum X_i \sum X_i = 5 * 160 - 0 = 800$$

$$(\mathbf{X}'\mathbf{X})^{-1} = \frac{1}{800} \begin{bmatrix} 160 & 0 \\ 0 & 5 \end{bmatrix} = \begin{bmatrix} 0.2 & 0 \\ 0 & 0.00625 \end{bmatrix}$$

$$(\mathbf{X}'_{2 \times n} \mathbf{Y}_{n \times 1})_{2 \times 1} = \begin{bmatrix} \sum Y_i \\ \sum X_i Y_i \end{bmatrix} = \begin{bmatrix} 1 & 1 & 1 & 1 & 1 \\ 8 & 4 & 0 & -4 & -8 \end{bmatrix} \begin{bmatrix} 7.8 \\ 9.0 \\ 10.2 \\ 11.0 \\ 11.7 \end{bmatrix} = \begin{bmatrix} 7.8+9+10.2+11+11.7 \\ 8*7.8+4*9+0*10.2-4*11-8*11.7 \end{bmatrix} = \begin{bmatrix} 49.7 \\ -39.2 \end{bmatrix}$$

$$\mathbf{B}_{2 \times 1} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} = \begin{bmatrix} 0.2 & 0 \\ 0 & 0.00625 \end{bmatrix} \begin{bmatrix} 49.7 \\ -39.2 \end{bmatrix} = \begin{bmatrix} 0.2 * 49.7 + 0 * -39.2 \\ 0 * 49.7 + 0.00625 * -39.2 \end{bmatrix} = \begin{bmatrix} 9.94 \\ -0.245 \end{bmatrix}$$

$$\hat{Y} = 9.940 - 0.245X$$

$$\hat{\mathbf{Y}}_{n \times 1} = \mathbf{X}_{n \times 2} \mathbf{B}_{2 \times 1} = \begin{bmatrix} 1 & 8 \\ 1 & 4 \\ 1 & 0 \\ 1 & -4 \\ 1 & -8 \end{bmatrix} \begin{bmatrix} 9.94 \\ -0.245 \end{bmatrix} = \begin{bmatrix} 9.94 - 0.245 * 8 \\ 9.94 - 0.245 * 4 \\ 9.94 - 0.245 * 0 \\ 9.94 + 0.245 * 4 \\ 9.94 + 0.245 * 8 \end{bmatrix} = \begin{bmatrix} 7.98 \\ 8.96 \\ 9.94 \\ 10.92 \\ 11.9 \end{bmatrix}$$

$$\mathbf{e}_{n \times 1} = \mathbf{Y}_{n \times 1} - \hat{\mathbf{Y}}_{n \times 1} = \begin{bmatrix} 7.8 \\ 9.0 \\ 10.2 \\ 11.0 \\ 11.7 \end{bmatrix} - \begin{bmatrix} 7.98 \\ 8.96 \\ 9.94 \\ 10.92 \\ 11.9 \end{bmatrix} = \begin{bmatrix} 7.8 - 7.98 \\ 9 - 8.96 \\ 10.2 - 9.94 \\ 11 - 10.92 \\ 11.7 - 11.9 \end{bmatrix} = \begin{bmatrix} -0.18 \\ 0.04 \\ 0.26 \\ 0.08 \\ -0.2 \end{bmatrix}$$

$$\mathbf{e}'_{1 \times n} \mathbf{e}_{n \times 1} = \left[\sum e_i^2 \right] = [-0.18 \quad 0.04 \quad 0.26 \quad 0.08 \quad -0.2] \begin{bmatrix} -0.18 \\ 0.04 \\ 0.26 \\ 0.08 \\ -0.2 \end{bmatrix} = [0.148]$$

$$MSE = \frac{0.148}{3} = 0.049333$$

$$\mathbf{V}(\mathbf{B}) = MSE(\mathbf{X}'\mathbf{X})^{-1} = 0.049333 \begin{bmatrix} 0.2 & 0 \\ 0 & 0.00625 \end{bmatrix} = \begin{bmatrix} 0.009867 & 0 \\ 0 & 0.000308 \end{bmatrix}$$

$$\mathbf{V} \begin{bmatrix} \widehat{\beta}_0 \\ \widehat{\beta}_1 \end{bmatrix} = \begin{bmatrix} Var(\widehat{\beta}_0) & cov(\widehat{\beta}_0, \widehat{\beta}_1) \\ cov(\widehat{\beta}_0, \widehat{\beta}_1) & Var(\widehat{\beta}_1) \end{bmatrix}$$

$$\mathbf{Y}'_{1 \times n} \mathbf{Y}_{n \times 1} = \left[\sum y_i^2 \right] = [7.8 \quad 9.0 \quad 10.2 \quad 11.0 \quad 11.7] \begin{bmatrix} 7.8 \\ 9.0 \\ 10.2 \\ 11.0 \\ 11.7 \end{bmatrix} = [503.77]$$

```
MTB > Copy C3 C1 m1
MTB > Print m1      Xn×2
```

Data Display

Matrix M1

```
1   8
1   4
1   0
1  -4
1  -8
```

X_{n×2}

```
MTB > tran m1 m2
MTB > print m2      X'2×n
```

Data Display

Matrix M2

```
1  1  1  1  1
8  4  0 -4 -8
```

$$\mathbf{X}'_{2 \times n}$$

MTB > mult m2 m1 m3

MTB > print m3 $(\mathbf{X}'\mathbf{X})_{2 \times 2}$

Data Display

Matrix M3 $(\mathbf{X}'\mathbf{X})_{2 \times 2} = \begin{bmatrix} n & \sum X_i \\ \sum X_i & \sum X_i^2 \end{bmatrix}$

```
5   0
0 160
```

MTB > inver m3 m4

MTB > print m4 $(\mathbf{X}'\mathbf{X})_{2 \times 2}^{-1}$

Data Display

Matrix M4

```
0.2  0.00000
0.0  0.00625
```

MTB > copy c2 m5

MTB > Print m5 $\mathbf{Y}_{n \times 1}$

Data Display

Matrix M5

7.8
9.0
10.2
11.0
11.7

$Y_{n \times 1}$

MTB > mult m2 m5 m6

MTB > print m6

$X'_{2 \times n} Y_{n \times 1} = X' Y_{2 \times 1}$

Data Display

Matrix M6

49.7
-39.2

MTB > mult m4 m6 m7

MTB > print m7

$(X'X)_{2 \times 2}^{-1} (X'Y)_{2 \times 1} = B_{2 \times 1}$

Data Display

Matrix M7

9.940
-0.245

$B_{2 \times 1}$

$$\hat{Y} = 9.940 - 0.245X$$

```
MTB > tran m5 m13
MTB > print m13
```

Data Display

Matrix M13

```
7.8  9  10.2  11  11.7
```

```
MTB > mult m13 m5 m14
```

```
Answer = 503.7700
```

```
MTB > mult m1 m7 m8       $\hat{Y}_{n \times 1} = X_{n \times 2} B_{2 \times 1}$ 
MTB > print m8
```

Data Display

Matrix M8

```
7.98
8.96
9.94
10.92
11.90
```

```
MTB > copy m8 c4
```

```
MTB > Let c5 = 'y'-C4
```

```
MTB > copy c5 m9 e
```

```
MTB > tran m9 m10
```

```
MTB > print m10 e'
```

Data Display

Matrix M10

```
-0.18  0.04  0.26  0.08  -0.2
```

```
MTB > mult m10 m9 m11 e'e
```

```
Answer = 0.1480
```

MSE=0.1480/3=0.049333 $MSE = \frac{e'e}{n-2}$

```
MTB > mult 0.049333 m4 m12
```

```
MTB > print m12
```

$$V(\mathbf{B}) = MSE(\mathbf{X}'\mathbf{X})^{-1}$$

Data Display

Matrix M12

```
0.0098666  0.0000000
```

```
0.0000000  0.0003083
```


$$V \begin{bmatrix} \widehat{\beta}_0 \\ \widehat{\beta}_1 \end{bmatrix} = \begin{bmatrix} \text{Var}(\widehat{\beta}_0) & \text{cov}(\widehat{\beta}_0, \widehat{\beta}_1) \\ \text{cov}(\widehat{\beta}_0, \widehat{\beta}_1) & \text{Var}(\widehat{\beta}_1) \end{bmatrix}$$

Regression Analysis: y versus x

Analysis of Variance

Source	DF	Adj SS	Adj MS	F-Value	P-Value
Regression	1	9.6040	9.60400	194.68	0.001
x	1	9.6040	9.60400	194.68	0.001
Error	3	0.1480	0.04933		
Total	4	9.7520			

Model Summary

S	R-sq	R-sq(adj)	R-sq(pred)
0.222111	98.48%	97.98%	94.11%

Coefficients

Term	Coef	SE Coef	T-Value	P-Value	VIF
Constant	9.9400	0.0993	100.07	0.000	
x	-0.2450	0.0176	-13.95	0.001	1.00

Regression Equation

$$y = 9.9400 - 0.2450 x$$

H.W

Q5.2 For the matrices below, obtain (1) $A + C$, (2) $A - C$, (3) $B' A$, (4) AC' , (5) $C' A$.

$$A = \begin{bmatrix} 2 & 1 \\ 3 & 5 \\ 5 & 7 \\ 4 & 8 \end{bmatrix} \quad B = \begin{bmatrix} 6 \\ 9 \\ 3 \\ 1 \end{bmatrix} \quad C = \begin{bmatrix} 3 & 8 \\ 8 & 6 \\ 5 & 1 \\ 2 & 4 \end{bmatrix}$$

Q5.5 Consumer finance. The data below show, for a consumer finance company operating in six cities, the number of competing loan companies operating in the city (X) and the number per thousand of the company's loans made in that city that are currently delinquent (Y):

i	1	2	3	4	5	6
X_i	4	1	2	3	3	4
Y_i	16	5	10	15	13	22

Assume that first-order regression model (2.1) is applicable. Using matrix methods, find (1) $Y'Y$, (2) $X'X$, (3) $X'Y$.